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MEMORY CODES REVEALED BY
RESPONSE LATENCIES

by

Lee Frederick DeCoster

A dissertation submitted in partial fulfillment
of the requirements for the degree of
Doctor of Philosophy
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DEDICATION

This work, as the culmination of my years in college and graduate school, is dedicated to my parents, Francis and Hilda DeCoster, in appreciation of the steadfast love and nurturance which they never failed to provide in times of celebration and times of despair; and to my sons, Barton Trent and Robb Frederick, as a source of encouragement that they also pursue those goals toward which they are inclined.

In keeping with this dedication, my writing has sought to minimize the use of technical jargon and the discussion of cryptic theoretical issues. A dissertation is a scholarly work. I believe the highest form of scholarship accepts the responsibility of providing succinct communication for the interested layman. To the best of my ability, I pursue that goal.

ACKNOWLEDGEMENTS

I gratefully acknowledge the time and effort contributed by all members of my dissertation committee. For personal reasons, I am especially pleased to have Drs. Melton and Pollack serve on this committee. Dr. Melton expressed faith in my potential as a psychologist at a time when I considered leaving this discipline. Dr. Pollack's confidence in my ability, patience with my speculations, and delicate guidance of my research and writing aspirations permitted the formulation and execution of this dissertation. Without the generous use of his facilities a project of this magnitude could not have been launched. Dr. Pollack's constant availability for the discussion of all aspects of this research will serve as a model for my future interactions with students.

Sylvan Kornblum first suggested an investigation of sequential categories of responding would be relevant to the serial exhaustive search model. The veracity of his request is apparent in the data.

Additionally, I thank Miss Barbara Friedman for programming, Miss Lynne Davis for data analysis, and Mrs. Nancy Mandell for the typing and proof-reading of an earlier draft.

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CHAPTER I

THE PROBLEM

Scientific endeavors can be partitioned into two types of activities: generating questions and developing methodologies. The inductive process of constructing theories and the deductive process of formulating predictions are inherent aspects of the former activity. When incompatible predictions are derived from separate theories the scientist has identified an issue. The resolution of issues requires sound methodology for the proper comparison of theories.

Psychology, as a science, inherits the task of questioning the nature of mental operations. That cognitive events are not observable, necessarily being inferred from measurable events, places a heavy burden on the inductive process. It is this extreme complexity of scientific inference involved in unveiling mental processes which prompts psychologists to explore and perfect new methods and experimental paradigms.

Recently, Saul Sternberg has developed both an experimental paradigm (Sternberg, 1966) and a rational-logical theory for operationally differentiating various cognitive processes (Sternberg, 1969; 1971). The present study attempts to apply these procedures and tools of inference, which were developed in another context, to a currently debated theoretical issue. The method examines response latencies in a memory-dependent, two-choice classification task. These time measures should be helpful in resolving the question: "What are the functional properties of those memory traces used in short-term recognition memory?" While the attributes of various types of memory have been investigated through a wide range of procedures, the form of the memory trace used in this type of task has not been determined. A clarification of the theoretical issue concerning memory precedes an elaboration of the methodology.

THE ISSUE

What Form Memory?

Introspection leads us to believe that words are stored in memory with large amounts of retrievable information. When asked, we can usually provide correct answers to questions about the spelling, pronunciation, overall shape, syntactical classification, and various aspects of the meaning of thousands of words. However, further reflection suggests that many things we could report accurately are not immediately available, but come after some special mental activities. For example, if asked how many letters are in the word "little," we usually need to first spell out the word and then count each letter. Moreover, we often have gaps in the usual array of memorized attributes of a word. Most of us need more than a second's time to recognize that "table" can be a verb, yet few adults require half that time to pronounce the word. Often a common word will not be immediately recognized when it appears in all capital letters, such as STIR, or POESY; conversely, what is the meaning of "i"? As a youngster I understood the meaning of "epitome" for years before I connected it with the spelled form, although I had seen it written many times. Many of us have a vocabulary of ethnic words which we understand and use, although we have never knowingly seen them in print. The word is in memory, but not in all of its forms.

While we may not wish to put much faith in such anecdotal evidence, we should not ignore its value in calling to our attention the fact that some properties of memory are immediately available while others require more time. Since we wish to know the "energy form" (Posner, 1967) or the functional code by which words are retained in memory, we must take a position regarding (1) the interpretation of variations in the time required to answer questions about the properties of words we have memorized, and (2) the type of memory we are prepared to investigate.

First let us consider that we observe a reliable difference in the time needed to correctly report one property of a word (e.g., its pronunciation) relative to some other property of the same word (e.g., its meaning).

We can interpret this finding in at least three ways: (1) There is a single memory code for the word, from which we can derive other features by performing additional cognitive operations. For example, if memory were auditory, we might pronounce the word, and then determine which letters were needed to produce such a sound by applying linguistic rules. We would then report the visual form of the word as calculated from its pronunciation. (2) There are several separate memory traces, all specifying different properties of the same word. However, it merely takes longer to either retrieve or to "decode" some types of these traces than other types. Thus we might have a visual memory trace and a set of semantic memories (synonyms), but we may be faster at locating the visual traces. (3) There are several memory traces for each word, all of them being chained together in such a way that we must search the chain in a fixed sequential order. For example, the chain might always first supply the spelling or visual form, then the pronunciation or auditory form, and finally the meanings of the word. While these three possibilities do not provide an exhaustive account of models of memory, they serve to give us a feel for the complexity of interpreting the hypothesized differences in response latencies. Throughout this thesis we shall employ the first model, which assumes (in an admittedly naive manner) that additional time to respond implies additional processing. The "more time, more processing" approach will be explained in greater detail in the later section concerning methodology. Although we might generate elaborate theories to deal with our topic, it will be to our advantage to construct the simplest possible model, modifying it only when the evidence is largely incompatible with the model's predictions.

The specification of the type of memory to be investigated is a more formidable task. In 1958, Donald Broadbent proposed that there were at least two structurally distinct types of memory: a short-term store which held information for time intervals "of the order of seconds," and a more permanent or long-term store. Since that time, there has been a proliferation of titles for different forms or different stages of memory operations. Extreme positions have been recently stated by Murdock (1970), who claims

there is no need to distinguish types of memory, and by Wickelgren (1970), who argues that there is empirical justification for at least four stages of memory. These stages have been differentiated with respect to such factors as their means of acquisition, the amount of information held in store, processes of consolidation and decay, and the operations required for retrieval.

Donald Norman (1968) argues that there is one stage, primary or short-term memory, which involves an auditory or acoustical representation, whereas a second stage, secondary or long-term memory, is thought to involve some type of semantic representation. The basis for this inference relied upon (1) an analysis of errors in recall or recognition using confusion matrices (e.g., Conrad, 1964, 1967), and (2) upon the amount of interference produced by stimulus materials varying in similarity with the to-be-memorized item (e.g., Dale & Gregory, 1966; Posner & Konick, 1966; Wickelgren, 1965). The similarity may be judged according to various dimensions, such as visual, articulatory, phonemic, or semantic similarity. Recently these two procedures have come into question due to their imprecision in discriminating among different forms of representations (Wickelgren, 1969), and due to their lack of agreement with results obtained using other methodologies (Shulman, 1970; Wickens, 1970).

We shall reverse Norman's approach. Rather than use evidence regarding memory codes to identify different types of memory, we shall leave open the theoretical specification of the type of memory under consideration in order to concentrate on how we might obtain evidence to clarify the nature of memory codes. Norman's "facts" become our issue. We shall take the objective position that we are studying whatever type of memory is required by our task. When we describe the task as a short-term recognition task, we mean by "short-term" only that the test comes less than a minute after presentation of the to-be-memorized words. By "recognition" we mean the subject is shown a set of items, some being the same as the memorized item, others being different. The subject must state (recognize) upon presentation of each item whether it was the same or different from the memorized item. It is in this sense that we ask, "What are the functional codes used in short-term recognition memory for verbal items?"

Three Hypotheses

The results of various lines of research enable us to formulate three hypothesized forms of the memory trace. In keeping with our emphasis on simplicity, we shall consider them as mutually exclusive possibilities within a single trace, although it should be clear that combinations of functional properties are possible (e.g., multi-dimensional trace theories have been proposed by Bower, 1967, and by Norman & Rumelhart, 1970). The Auditory Hypothesis proposes that a memory trace preserves only the sound features of a word, being isomorphic with the perception of spoken words. The Semantic Hypothesis claims that a memory trace retains the meaning of a word. Finally, the Visual Hypothesis states that a memory trace is isomorphic with the physical-spatial form of written words. Each hypothesis will consider that a word represents an auditory, semantic, or visual unit, although each unit may have subunits or features. Sample evidence related to each hypothesis will be considered in order.

The principal evidence for the Auditory Hypothesis has come from error analyses. Confusion matrices can be understood through comparisons between different letter pairs. Consider that we briefly flash the letter "B" on a screen and a subject erroneously reports the letter "P." We are not at liberty to determine whether this error resulted from the similarity in the visual form of the two letters, the similarity in their sound, or their common means of articulation. However, confusing an "E" for an "F" might incline us to attribute the error to visual similarity, since the letters are not articulated alike, nor do they have the same physical sound pattern. By the same reasoning, confusing a "B" with any of the following letters -- C, D, G, T, V, or Z -- would be attributed to the acoustic similarity (same trailing phoneme) rather than visual similarity. This implies that the letters had an auditory code and not a visual code.

Wayne Wickelgren (1969) pointed to the difficulty in differentiating among acoustic hypotheses by asking, does the subject encode the way an item sounds (auditory) or does he store that set of muscle commands required to pronounce the item (articulatory form)? Unfortunately, these

two possibilities, while conceptually distinct, are naturally confounded for English letter names. Although unmentioned by Wickelgren, we should further note that the physical wave form of sound is correlated with auditory perception. In general, letters which are perceived as distinct have dissimilar patterns of speech production and have greater physical differences in spectral frequencies. As an example of similarity in perception and articulation, we can consider the pairs B-D, M-N, and D-T. These pairs sound similar (perception) and have similar articulatory features. The articulatory difference between B and D, and between M and N is simply a difference in the "place" of articulation, while the difference between D and T involves a difference in "voicing." In general, the more dimensions of articulation which differ, the greater the difference in both perception and wave form. To avoid theoretical issues of speech perception, we will avoid differentiating these hypotheses, using the terms auditory code and acoustic form of memory to refer to the same general hypothesis, their interpretation being that the memory trace is isomorphic with the perception of a spoken word.

Interference studies typically give subjects a list of words and provide a test of recognition, or recall, immediately thereafter. By constructing lists containing homonyms, synonyms, or unrelated words, comparisons of accuracy in recall are possible. For example, when this procedure was used by Kintsch and Buschke (1969), using 16-word lists followed by single-word cued recall, the authors found less accurate recall of early items in the synonym list, relative to a control. Accuracy was less for later items with the list containing homonyms. These performance decrements relative to a control were attributed to interference, presumably produced by similarity between memory traces. These data suggest that the memory of words which have been available for longer time intervals contain attributes of meaning, while memory traces available for shorter time periods are primarily auditory.

Another procedure involving interference provides further evidence of semantic properties in short-term memory. Delos Wickens (1970) reports that recall of word triads was near perfect for the very first trial, but that accuracy dropped considerably on the second and

third trials. By the fourth trial performance was very near an asymptote (plateau) of less than 60% correct recall. Although additional trials with the same types of words continued to show this low level of accuracy, it was possible to have the subject's performance return to its initial, near-perfect level. This was done by changing the stimulus materials. Wickens argues that the demonstration of a return to a high level of performance implies that any drop in accuracy must have been due to interference. Once again, it is assumed that this proactive inhibition was produced by similarity between memory traces. Changing the critical stimulus dimension means the new materials are functionally dissimilar, hence the source of interference has been removed. By this reasoning, those dimensions which, when altered, fail to eliminate interference must not be represented in the memory trace. They are irrelevant to the memory code. Conversely, release from proactive inhibition signifies that the altered dimension is a critical component of the trace. Wickens found that changing physical features, such as the color or size of the words had little effect. Altering the taxonomic category of the words, as in going from names of flowers to names of trees, substantially improved performance. Therefore these semantic categories must be represented in the memory code. Consequently, Wickens claims evidence for the Semantic Hypothesis.

Additional evidence for semantic encoding has been supplied by Harvey Shulman (1970) based on an analysis of reaction times. However, since this procedure is an outgrowth of previous reaction time research conducted by Posner and by Cohen, it will be instructive to digress from the chain of evidence supporting the Semantic Hypothesis in order to consider the earlier research.

Concern with the temporal properties of word and letter recognition springs largely from experiments using Sperling's (1960) procedure wherein large arrays of letters are presented visually for a brief duration followed by a cue to recall a subset of the array. Such studies have provided evidence of confusions related to the visual form of the stimuli (Chase & Posner, 1965; Keele & Chase, 1967). Sperling (1963, 1967) interprets the difference in the types of errors resulting from this

task, relative to others using longer exposures and delayed recall, as evidence that there is a visual form of memory which holds information until it can be transformed into an auditory representation.

Support for this interpretation has come from reaction time experiments by Michael Posner and his colleagues (Posner & Keele, 1967; Posner & Mitchell, 1967; Posner, Boies, Eichelman, & Taylor, 1969). The procedure in these experiments involved the sequential presentation of a pair of letters, only one letter being visible at a time. A decision was required after each pair. Typically, the first letter was presented to the subject in either upper case (e.g., "A") or lower case print ("a"). The instructions were to respond "Same" if the second test letter had the same name as the first letter, otherwise respond "Different." The term "name" referred to the way a letter was pronounced. The reaction time was quicker when the pair of items were physically identical, relative to being similar only in name. The advantage of the identity match was greatest when the interval between the removal of the stimulus standard and the onset of the test letter was short. The non-identical name match was at least as rapid for inter-stimulus intervals greater than 1.5-2.0 seconds. Thus the Visual Hypothesis finds support in studies where the memory trace must be used just briefly after it has been established.

Gillian Cohen (1968) made use of another of Posner's procedures -- the simultaneous presentation of letter pairs. Cohen used word pairs, rather than sets of letters. The pairs were either identical (sew-sew), homonyms (sew-so), or synonyms (sew-stitch). Each subject made a Same vs. Different judgment to each pair by depressing response keys using separate hands. The basis for the response classification was one of three instructional rules, all subjects working with each instruction. These rules required a Same judgment if the words were (1) visually similar, (2) acoustically similar, or (3) semantically similar, as in the above examples.

Perhaps the most interesting result involves the reaction time to identity pairs when the various instructions were in effect. The Same response required 678 msec. under visual match instructions, 716

related to list items. The subject was to respond Yes (i.e. the item is equivalent to one in the list) or No to the probe.

Shulman had hoped that the rule uncertainty, produced by delaying the instruction until after the list had been presented, would force semantic encoding to the degree that such encoding was possible. He further reasoned that semantic encoding (setting up a memory trace with semantic properties) may require more time to establish relative to an auditory trace. Consequently, the rate of word presentation was varied, having single-item exposure durations of 250, 600, or 1300 msec. Shulman predicted that the response time to synonyms would be shorter with slower rates of presentation.

The results showed equal proportions of correct Yes responses for identity items and homonyms, but fewer correct synonym judgments. Reaction times were quickest for identity items, next for homonyms, and slowest for synonyms. All three functions produced similar patterns of response times as a function of the position in the list of the matched item. The average difference in Yes responses was about 165 msec. between identity items and homonyms, and approximately 255 msec. between homonyms and synonyms. While the reaction times increased for all types of items with slower rates of presentation (contrary to the prediction for synonyms) the proportion of false positives decreased for synonyms with slower presentations while increasing for homonyms. Although Shulman interpreted both reaction time and error data as evidence for semantic encoding, it is only this last finding which can not be explained solely from the Acoustic Hypothesis (assuming the sound is retrieved and the meaning then derived from some long-term store). However, since the error rate for synonyms was always greater than for homonyms, even this interaction of errors by rate of presentation is unconvincing. Additionally, we should note that greater speed and fewer errors for probes identical to test items supports the Visual Hypothesis.

The issue remains: Is short-term memory essentially visual, auditory, or semantic? The bulk of evidence supports the Auditory Hypothesis, but this model is unable to account for faster responding to visually-similar materials. Those supporting the Visual Hypothesis

Donders proposed that the magnitude of the difference in response times (comparing the two situations) represents the time needed for additional stages of processing. This Subtraction Method assumes "pure insertion" of a new stage of processing. That is, the slower reaction times which result from increased task requirements are produced by the addition of functioning centers, rather than increments in the time needed at each existing stage of processing. Technically, these stages must be stochastically independent.

Sternberg (1971) argues that the strong theoretical assumptions of the Subtraction Method are not necessary. He proposes an Additive-Factors Method that also requires the assumption of sequential processing, but provides that an experimental variable can affect more than one stage. It is possible to determine which experimental manipulations affect which inferred stages of mental processing by examining the data. Data need to be statistically evaluated for main effects and interactions in order to make the necessary inferences. This model does not suggest that processing centers operate in an all-or-none fashion. Even with these less restricting assumptions, the method is capable of establishing both the number of separate functional (inferred) stages which are operating, and the time parameters for any given stage. A more complete account of this theoretical framework for interpreting reaction time data can be found in Sternberg's works (1969, 1971).

The Experimental Paradigm

Sternberg (1966) also developed a particular experimental paradigm which helps specify a crucial feature of memory (viz., a search to determine the presence or absence of a code). The procedure calls for presenting a list of from one to six items to a subject for memorization. The subject can be given a few seconds or several minutes to memorize the items which can be presented either sequentially or simultaneously. Then a single-item test is displayed visually. The subject's task is to either press "Yes," signifying that the item is the same as one of the memorized items, or "No," that it is different. Next the procedure may call for

either a new list of to-be-memorized items (thereby having only one test per trial), or several more tests can be given for the same memorized items. The to-be-memorized items are said to form the Positive Set since they require a Yes response. The dependent variable of greatest interest has been the response latency, the time measured from the onset of the test item to the pressing of a response key. Separate hands are used to indicate a Yes vs. No response.

Typical results show a mean response time of just over 400 msec. for both Yes and No responses when a single digit or letter has been memorized. The Yes time is typically faster when Yes and No responses are equally likely. Increasing the probability of either response will lower its average latency relative to the alternative response. The average reaction time increases in equal increments of 35-40 msec. as the number of items in the Positive Set increases. Consequently, the easiest way to describe the data is to plot reaction time as a function of the size of the Positive Set, in which case a straight line having a slope of 35-40 msec. per item connects all observations. If the lines are fitted separately to all Yes and all No responses, the lines will be nearly parallel, the No line intercepting the ordinate slightly above the Yes line. The interpretation of such data is quite straightforward, although the theoretical argument is detailed. It appears that a subject recognizes the test item and then compares this item with each-and-every item in memory. A Yes response is initiated if the test item matched one of the memorized items; a No response being made in the absence of a match. Thus the subject makes a serial, exhaustive search of the Positive Set.

The theoretical argument considers three questions, each issue having two possibilities. They are: (1) Are the comparisons between the test item and the items in memory made one-at-a-time (serially) or simultaneously (in parallel)? (2) On correct Yes tests, is the search ended as soon as a match is made (self-terminating) or are all possible comparisons made before the response is initiated (exhaustive)? (3) Is the length of time to make a single comparison a constant, or is it variable (zero vs. non-zero variance)? A fourth, contingent issue is:

When there is more than one test after a list has been memorized, there are possibilities of both response and stimulus repetitions. Consider that there are two items in the Positive Set, the digits "1" and "2." Presenting a "3" as the first test item results in a No response. A "2" as the second test item requires a Yes response. Since this second response is a different one, it will be termed a Non-Repetition. The change could have gone from a Yes first response to a No, in which case we would classify the second response as a No Non-Repetition. If on the third test a "1" is presented and the subject responds Yes again, we have a response repetition. However, the stimulus has changed, constituting a stimulus non-repetition. Following Bertelson (1965), this form of response repetition will be termed an Equivalence response. Presenting another "1" as the fourth test would involve both a response repetition and a stimulus repetition, which we will call an Identity response. It is possible to have Identity or Equivalence responses for both Yes and No classifications provided the Positive and Negative Sets contain more than one item (with one item a Yes response repetition must be an Identity).

If a subject always searches through the entire Positive Set before initiating a response, there is no reason to expect a difference in the reaction time among the repetition categories. However, Sylvan Kornblum (1969) has shown that sequential effects can be an important determinant of response latencies, affecting a variety of tasks. Repetition responses require less time than non-repetitions, and consecutive repetitions show a continual reduction in latencies. What would be the interpretation of a main effect of repetitions in the Sternberg paradigm? This is where the Additive-Factors Method is needed. The method states that an experimental variable producing a main effect, but no interaction with other variables, influences a separate stage of processing relative to the other variables. Thus if repetitions have a lower intercept than non-repetitions, but have the same slope as a function of the size of the Positive Set, it is safe to assume that this variable affects some stage of processing other than the memory-comparison stage. Presumably the effects would be due to a facilitation in some process which controls the muscle action -- a response process.

A simplified statement of Wattenbarger's rationale attends to the difference between the slopes for the Physical Identity vs. Name Identity Conditions. If memory involves visual images, twice as many comparisons would be necessary for the Name Identity Condition relative to the Physical Identity Condition since, in the former, subjects would need to store both the upper and lower case version of each letter. Consequently, the slope for the Name Identity Condition, when plotted as the number of distinct letter names, would be twice as steep as that of the Physical Identity Condition, where the slope is plotted as the number of physically distinct visual patterns. On the other hand, if memory codes are auditory, the shallowest slope will be in the Name Identity Condition. In the Physical Identity Condition the subject will presumably store both the letter name and some additional code denoting which case (upper or lower) is correct. If this additional code requires a comparison time equal to that for the letter name, the slope of the Physical Identity Condition will simply be twice as steep as that of the Name Identity Condition.

The obtained slopes, combining Yes and No responses, were 97 msec. per character for Physical Identity vs. 57 msec. per letter name for Name Identity. Based on the above reasoning, these data support the position that memory codes are names, or in the terminology of the present study, auditory codes. The Control Condition slopes were 54 and 44 msec. per letter for upper and lower case characters, respectively. The similarity of these slopes to that of the Name Identity Condition is consistent with the conclusion that name or auditory codes were used throughout. However, contrary to our Auditory Hypothesis, a breakdown of the No responses in the Physical Identity Condition revealed a steeper slope (102 vs. 89 msec. per character) and higher intercept (816 vs. 555 msec.) for those items having the same name but a different case, relative to items differing in both name and case. Finally, the error rate was lower for the Name Identity Condition relative to the Physical Identity Condition, 3.1% vs. 6.0%, again supporting the Auditory Hypothesis.

The data from Wattenbarger's study are less than satisfying for two reasons: (1) the range of response times, for both slopes and intercepts, departs from other studies, and (2) the data provide no consistent

clue as to what the subject was doing in the Physical Identity Condition. The 57-msec. slope of the Name Identity Condition is considerably greater than the 35- to 40-msec. slopes obtained by Sternberg (1969) using digits. Also, Sternberg's intercepts were typically below 400 msec., compared to the 528- and 630-msec. intercepts of Wattenbarger's two principal conditions. The Auditory Hypothesis, as applied to Wattenbarger's design, predicts (1) the intercepts for the Physical Identity and Name Identity Conditions should be identical, and (2) there should be no difference between slopes or intercepts for the two types of No response in the Physical Identity Condition. Neither prediction was confirmed.

An alternative explanation of Wattenbarger's findings assumes the Physical Identity Condition elicited a different form of memory trace (viz., visual) which required more time for single comparisons. But this model also makes the unverified prediction of no difference between the two types of items requiring a No response (differ by case alone vs. differ by both case and name). At best, Wattenbarger's results suggest that a memory search process can be conducted most efficiently through auditory memory traces. Whether or not other forms of memory trace were used can not be determined.

ORGANIZATION OF UPCOMING CHAPTERS

Now that the issue and methodology which prompt this thesis have been outlined, we are prepared to consider the procedures and results of the present research. Six chapters describe the procedures, report the results, and discuss the implications of three experiments. Chapter II provides (1) an overview of the experimental design, (2) an account of the assignment of subjects to conditions, and (3) a description of the equipment and procedures of Experiment I. Since the procedures of the other two experiments were highly similar to those of the first experiment, any modifications will be explained as each additional experiment is introduced. Chapter III reports the findings of Experiment I for both response latencies and errors. The fourth chapter discusses puzzling results of Experiment I, describes the rationale and results from Experi-

ments II and III, and outlines those aspects of the data requiring clarification. Chapter V resolves most of the remaining complications by partitioning the data according to sequential patterns of responding. A final analysis of the results from Experiment I takes place in Chapter VI, preceding the conclusions of the final chapter. In general, the presentation of results and the accompanying discussions are organized about a single theme -- what is the functional code used in short-term recognition memory?

CHAPTER II

OVERVIEW OF EXPERIMENTATION

Design

Three experiments were needed to satisfactorily resolve the basic issue. Experiments I and II had a common design. Both involved the use of three experimental conditions, all subjects serving in every condition. All possible orders of serving in the three conditions could be incorporated by using six subjects, each receiving a different sequence of instructions (i.e., 1-2-3, 1-3-2, 2-1-3, etc.). Experiment I used one six-subject block whereas Experiment II used two such blocks (12 subjects).

Each of the three conditions required a one-hour testing session. The term series refers to a set of three sessions wherein each of the three conditions was represented once. Experiment II used a single series, whereas Experiment I involved three series. That is, each of the six subjects in Experiment I participated in each of the three conditions three times, amounting to a total of nine one-hour testing sessions (these subjects also had a one-hour practice session, bringing the total to 10 hours). No subject worked more than one session on a given day.

Different subjects were used for each experiment, except that three subjects from Experiment I served in Experiment II for four sessions at a single condition. Their participation was an addition to the above-described design, incorporated to obtain a more stable estimate of performance in this condition. Also, an additional subject for each condition worked for a single session in Experiment II. This was done in order to provide a more reliable estimate of first session performance since these responses would be uncontaminated by any possible influence of

serving in other conditions. Thus a total of five subjects worked in each condition as their first session in Experiment II.

Experiment III involved six subjects serving one session in the only condition in that study. One subject was subsequently eliminated, reducing the total to five subjects. Since this condition called for subjects to attempt to eliminate the naming of words, it seemed unwise to have these subjects participate in those conditions which required that the to-be-memorized list be read aloud and actively rehearsed.

Subjects

A total of 18 females and 11 males, all enrolled at the University of Michigan, began as subjects, with the data from 17 females and 9 males being reported. The female:male breakdown by experiment was 5:1, 10:5, and 2:3 for Experiments I, II, and III, respectively. All but two subjects were right-handed, with one additional subject reporting he was ambidextrous, although he wrote consistently with his right hand.

The subjects in Experiment I were all experienced, working more than four hours per week in a computerized psychophysics laboratory. Each was paid \$2 per hour, plus a bonus of \$2 for completing the 10 sessions. The first session for each consisted of instruction and practice at all three conditions of the experiment. Thereafter these subjects set up the equipment and worked alone.

Naive subjects served their first hour in Experiment II to fulfill a psychology course requirement. After completing that hour, each was asked to return for two additional sessions at \$2 per session. All but one of the subjects volunteered to return for pay. The data from the non-volunteer (who expressed the wish to return, but an inability to spare the time with several exams scheduled at the time for experimentation), and that from two other participants filled the need for three one-session-only subjects. The experimenter was present only during the seven practice trials in each session during Experiment II. The subjects worked alone for the test trials.

Subjects in Experiment III were obtained from the Psychology Department Subject Pool in the same manner and at the same time as those in Experiment II, but these subjects were not asked to return for pay. Those selected for this experiment were in no apparent way different from those serving in Experiment II. The order of assignment to each experiment had been determined in advance. However, one subject produced such long latencies (the median being over one second), and expressed such a negative attitude toward the research, that her data was discarded prior to analysis.

In addition to the above subject, two subjects in Experiment I were dismissed -- the one being unable to serve the total of 10 sessions (one practice, nine testing sessions), the other being asked to stop when he was unable to reduce his error rate below 5% after three sessions. Partial analysis of the data from both subjects revealed no major discrepancies from the results of the other subjects. Nevertheless, the a priori criterion of dealing with low error rate subjects was administered.

Equipment and Parameters

The stimulus arrays, stimulus parameters, and response measurements were constant throughout all experimentation. The visual display was presented on a modified Tektronix Model 611 storage display oscilloscope, controlled by a Digital Equipment Corporation Model PDP-8I computer. The letters were all capitals plotted upon a potential 9 x 13 grid with letter size being approximately 3.7 x 5.4 mm. With the subject seated approximately one meter from the scope, a three-letter word would subtend 0.8° of arc.

The subject indicated his decision by pressing one of two micro-switch response keys, mounted on a table in front of the scope. A 5- to 6-ounce force was sufficient to trigger the switch, depressing the key 1/8 inch. The subject was free to use either his index or second ("middle") finger, one finger for each hand. The hand used in writing was designated for the "Yes" response. All responses were timed on the computer's counter, accurate to ±1 millisecond.

A trial was defined as the sequence of (1) an instruction, (2) a to-be-memorized list, and (3) a series of test items, presented one at a time. The instruction was always a two-line display, reading "INSTRX:" on the first line, and "SOUNDS THE SAME," or whatever instruction defined the condition, for the second line. The instruction remained for two seconds, followed by a one-second blank period, and then the stimulus array. The to-be-memorized list of stimuli was arranged vertically, the top-most item centered on the screen. The remaining items were listed below, centered according to the rule of adjusting one place to the left for every two letters present (thus one- and two-letter items started at location zero, three- and four-letter words starting at position -1, five- and six-letter words at position -2). The duration of this display was $1667 \text{ msec.} + (d)333 \text{ msec.}$, where d = the number of displayed stimuli. A three-second blank period followed, before a test item appeared, always centered. The test item remained on the screen until the subject responded. However, the time was not recorded if it extended beyond two seconds (this was extremely rare, even during the practice sessions).

Upon a response, a two-second blank period elapsed before the next test item appeared. Trials for which data were evaluated always involved 10 tests per trial, but practice trials had as few as four tests per trial, varying between experiments. Two seconds was the minimum time following the end of one trial before the instructions appeared for the next trial. This time interval varied, since a slow tape reader had to supply a variable length list of parameters and stimuli for the upcoming trial following the last response to the previous trial.

While the stimuli for a given trial were determined by a parameter tape, the sequence and content of test items was controlled by the computer via the following rules: There were always two permanent word pools: (1) the 12 potential stimuli, any of which might form the Positive Set of memorized stimuli on a given trial, and (2) a Null Set of 12 items which could never appear as stimuli, and hence never could be correctly classified with a Yes response. The latter items did appear as tests, always requiring a No response.

Figure 1 illustrates both the oscilloscope displays and the computer's organization of word pools. The probability of displaying a test item which belonged to the Positive Set was always 0.50. These items required a Yes response. A random selection determined (1) which word pool was consulted and (2) given a pool, which item was to be displayed as a test. All selections were made with replacement and the probabilities had no string-length constraints. A new Positive Set was defined for each trial based on the instruction and the to-be-memorized list, as shown in Figure 1. Once again, the latter information came from a parameter tape which was read before each trial began. We should also note that while the probability of a Yes response was constant, the probability of displaying a particular stimulus was confounded with set size (the larger the set size the less likely any one stimulus would appear).

On those tests requiring a No response, one of three categories of items could be selected. Once again a random number chart was used by the computer to determine which of two pools of items should be consulted. With probability 1/2 (overall $p = 0.25$), a test item was selected from a pool of Inclusive items, otherwise it came from the larger pool consisting of both Exclusive items and the 12 Null Set items. On those trials where there were no Inclusive items, the selection would automatically be from the second pool (Exclusive plus Null items).

The term Inclusive is appropriate for two conditions in Experiment I -- Looks and Means. An Inclusive item is defined as a homonym of a displayed stimulus when that homonym version requires a No response on the trial in question. If, as in Figure 1, the stimuli were the items "WON, 2," the Inclusive Set would contain the items "1, ONE, TWO, TOO" for the Looks Condition, but only "1, ONE, TOO" for the Means, since "TWO" means the same as "2" and consequently would be a member of the Positive Set (requiring a Yes response).

Exclusive items are those remaining from the 12 potential stimuli after all Positive Set items (correct for the trial in question) and their homonyms (Inclusive items) have been removed. In the illustrated

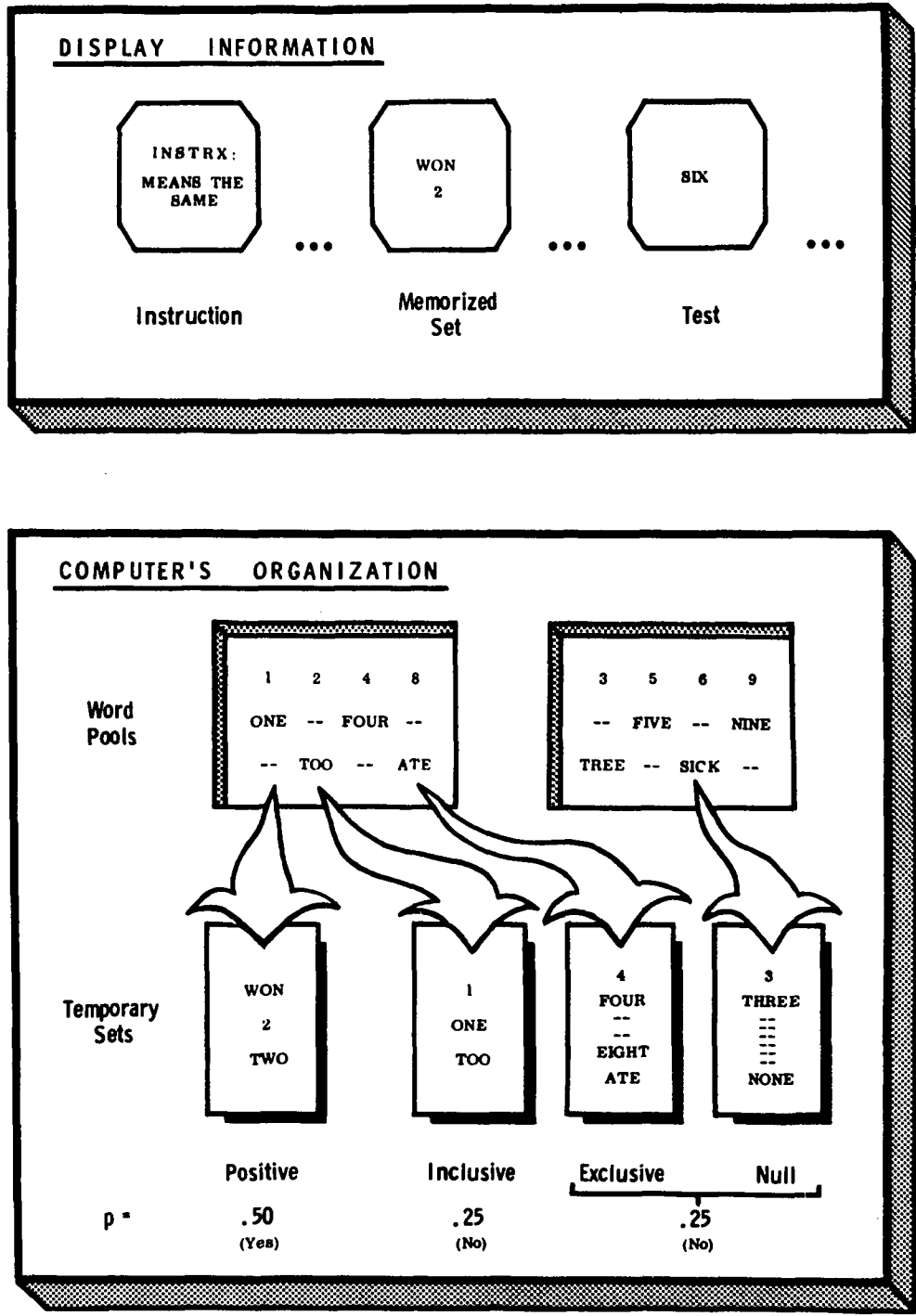


Fig. 1. Illustration of a trial, showing displayed information and the computer's arrangement of permanent word pools into temporary functional sets for the selection of test items.

(2) the number of potentially correct items = C, and (3) the number of "Types" represented in the display = T. The term "Types" refers to the categories: Arabic numbers, English spelled numbers, and non-number English words. Variables D and T are determined by the to-be-memorized list as it is displayed, and consequently these variables are not affected by the instruction. The variable C is determined by the instruction as it relates to the to-be-memorized list of stimuli. Table 1 illustrates these three variables in conjunction with the Sounds and Means instructions. For the Looks instruction, $C = D$. For all three instructions, D varies from 1-4 while T varies from 1-3. C takes the values 1-4 for the Looks Condition, 3, 6, 9, and 12 for the Sounds Condition, and 1-8 for the Means Condition.

In part, the instruction serves to define the D:C ratio. Since for the Looks instruction only the displayed items are correct, the D:C ratio is constant at 1:1. For the Means instruction the ratio is 1:1 only when the non-number words are displayed (see first two examples of Table 1). For the numerical stimuli, there is a 1:2 ratio between D and C (unless both forms of the number are displayed as stimuli, as in the third and fourth examples in Table 1). Thus displaying either numerical form, as in the first example of Table 1, signifies that both forms are correct, since they have equivalent meanings. Similarly, for the Sounds Condition, all members of a homonym triplet are correct when any one is displayed, resulting in a maximal 1:3 ratio (as in the first, second, and last examples of Table 1).

The trial sequence for Experiment I, which contains only a sample of all possible stimulus displays, was identical for all three instructions. [The contents of the trials are listed in Appendix A by trial number, whereas the sequence of trials is preserved in Appendices B, C, and D.] The sequence was arranged to maximize critical comparisons between the instructional conditions. The 48 data-collection trials were first divided equally among the four values of variable D (i.e., 12 trials displayed one item, 12 more displayed two items, etc.). For each value of D, the 12 trials were evenly divided among the possible values of C as defined by the Sounds instruction. For example, of the 12 trials where

three items were displayed, four displayed all three members of a homonym triplet such that $C = 3$, four displayed two sound classes such that $C = 6$, and the remaining four trials displayed three different sounds such that $C = 9$ (third, fourth, and fifth examples, respectively, from Table 1). Finally, within a given D-C pair with C calculated for the Sounds instruction, the Types of items were varied to include different values of C as calculated for the Means Condition. For example, with three displayed items incorporating two sounds, C values of three (WON, 8, EIGHT), four (TWO, TOO, FOR), and five (ONE, WON, FOUR) were represented in the Means Condition.

A difficulty with the above priority system is that while it equalizes the number of responses for each value of D (and therefore of C) within the Looks Condition, it creates an imbalance for the other two conditions, having more cases with small values of C and very few with C at its maximum. [This problem was adjusted in Experiment II by counterbalancing first on the variable C , and then on D and T , thereby equalizing the number of responses among the sets defined in auditory units.]

To the limit of the total number of displayed items, all stimuli occurred equally often in the to-be-memorized display. The order of trials was determined from a random number table with the constraint that no three consecutive trials contain the same displayed stimulus. [A counterbalancing method of trial ordering was used in Experiment II to distribute more evenly trials throughout a session.]

CHAPTER III

THREE CRITERIA

The results which best relate to our original question are found in Experiment I. It is here that the three competing hypotheses will be compared using identical materials but different instructions. Each instruction was designed to be most compatible with a single hypothesis. Presumably the Looks Condition calls to the subject's attention the visual features of the stimuli. Similarly, the Means Condition forces the subject to use the semantic features of the stimuli while the Sounds Condition requires attention to the auditory properties of memorized items.

Our first interest is in the degree to which the different instructions facilitated the speed of responding. But the average latency of responding does not provide a strong test of our hypotheses. The use of the Sternberg paradigm allows us to compare slopes between conditions. When reaction times are grouped according to the size of the Theoretical Positive Set, we can derive the average time to make a single memory comparison (i.e., the slope). Our three hypotheses suggest the form of memory used for this memory-matching process. By making the naive assumption that subjects search a visual memory in the Looks Condition, a semantic memory in the Means Condition, and an auditory memory in the Sounds Condition, we can interpret the slope of the Looks Condition as the estimate of the rate of search through a visual store, the slope of the Means Condition as the rate of searching a semantic store, and the Sounds Condition slope as reflecting the comparison time with auditory memory codes. In each case these slopes are calculated using the Theoretical Positive Set which is compatible with the instruction (i.e., Looks = Visual Hypothesis; Means = Semantic Hypothesis; Sounds = Auditory Hypothesis).

While it is conceivable that there are three distinct forms of memory, such that all three hypotheses are verified, the Law of Parsimony recommends we ask: "What pattern of results would indicate the presence

of a single form of memory?" We can deduce that a single memory code is being used if we can meet the following three criteria.

I. One condition results in a shallower slope than the other two and the reaction times are at least as fast.

II. There is less variability among trials within each Positive Set size in the favored condition of Criterion I relative to the intra-set variability of the other two conditions.

III. The hypothesis compatible with the favored condition (by the above two criteria) is capable of reorganizing the trials of the non-favored conditions according to new Theoretical Positive Sets such that the resulting data meet two standards: (1) the newly calculated slope is equal to that of the favored condition, and (2) the variability among trials within a set has been reduced relative to the original calculation of intra-set variability. A further discussion of these criteria follows.

Although there is an instructional condition which is most compatible with each hypothesis, we can also apply the other "less-compatible" hypotheses to each condition. For each instructional condition we can order trials according to the Visual, Semantic, or Auditory Theoretical Positive Sets. One way to conceptualize this 3 x 3 set of possibilities is by an analogy to a deck of playing cards. Let us say that we have been dealt a hand of seven cards. Without substituting any cards we order this hand by first arranging them by number value, then by suit. Alternatively, we could have first arranged the cards by suit, and then within each suit by number value. The same cards can be ordered two ways depending on the game we wish to play, or on the highest point value within a game. In a similar way, we have a set of trials which can be organized by each of the three hypotheses (redefining the Theoretical Positive Set for calculating the slope) separately for each of the three instructions.

We can get a feel for the validity of the first criterion through a brief example. First let us assume that there is an ideal observer who has a true intercept of 400 msec. (the sum time for all processing in the absence of a memory comparison process). Our observer takes 50 msec. to make a single memory comparison. Thus when he has one stimulus code in memory his response time to a test is $400 + 50 = 450$ msec. For set

and Y incorporate two auditory units while trial Z includes three auditory units. Let us further assume that memory is auditory. Consequently the response times for X and Y are nearly identical, but the time for Z is longer (by an amount equal to the time to search its one additional auditory unit). Grouping these three trials together under the Semantic Hypothesis creates unnecessary variability among trials of set size three. The range of response times is wide. However, when we reorganize the data of the Means Condition according to the Auditory Hypothesis, we correctly classify trials X and Y together as involving two sound units, while trial Z will be grouped with any other trials involving three sound units. The new Theoretical Positive Sets, defined in auditory units, show little variability among response times of similarly grouped trials. Moreover, we can calculate the slope of a line through the mean value for each set size. As in the previous example, we expect this slope to be equal to the slope of the Sounds Condition (calculated in auditory units).

Now that the appropriate means of evaluating the data has been outlined, we are prepared to launch into the results. First we shall consider the reaction time findings for Experiment I, to be followed by an analysis of errors.

EXPERIMENT I

Criterion I

The strategy for analysing the results of Experiment I will be to apply the three criteria to the data in order to identify a single form of memory, if such a conclusion is warranted. Only if this strategy fails will we conclude that more than one form of memory code is functional. As a preliminary step, the average weighted Yes and No response time was calculated for each trial, pooling the reaction times of the six subjects. Appendix A lists the items displayed on each trial while Appendices B, C, and D list unweighted trial means by series for the Looks, Means, and Sounds Conditions, respectively. [The weighted means preserve the information of every response whereas the unweighted means retain the average

time for each subject. In general, the weighted means will yield faster reaction times.] A highly summarized version of the data is presented in Table 2, where Yes and No response times have been pooled. These data permit application of Criteria I and II. Considering the absolute magnitude of the reaction times, the overall mean times (based on 8,640 timed observations) were 461, 494, and 513 msec. for the Sounds, Looks, and Means Conditions respectively. An analysis of variance test verified that the Sounds Condition was significantly faster than the other two conditions ($p < .01$ between Sounds and Looks). On the basis of the overall mean times, the criterion of lowest response times favors the Auditory Hypothesis.

Inspection of response times as a function of the size of the Theoretical Positive Set (Table 2) reveals that none of the conditions produced a linear slope through all four sets. However, the slope for the Looks and Sounds Conditions are nearly linear through three of the four points. This finding is more apparent in Figure 2 which displays results of these two conditions by series. The slope for the Sounds Condition (right side of figure) is clearly flatter, averaging approximately 30 msec./sound for sets 1-3 (the line accounting for 99.8% of the variance). The drop in latency for set size four is perplexing. Also confusing is the finding that the reaction time to a Theoretical Positive Set of one unit is faster for the Looks Condition than for the Sounds Condition. If memory is solely auditory, the latter should be at least as fast. We shall return later to these two puzzling results.

The slope of the Looks Condition (left side of Figure 2) attenuates from 54 msec./item to 45 msec./item with successive series. Unlike the value in Table 2, the plotted point for set size three does not include data from the four trials where all three versions of the same sound were displayed (i.e., a complete homonym triplet). These data will be discussed later. The curve for Series One is a near-perfect straight line when the abscissa is plotted logarithmically (similar to the data of Swanson & Briggs, 1969). However, the curve flattens to a straight line by Series Three. Series One's deviation from a linear function is related to the Yes response to a single item, where the observed latency is 60 msec. faster than that predicted from a straight line passed through the Positive Set points 2, 3,

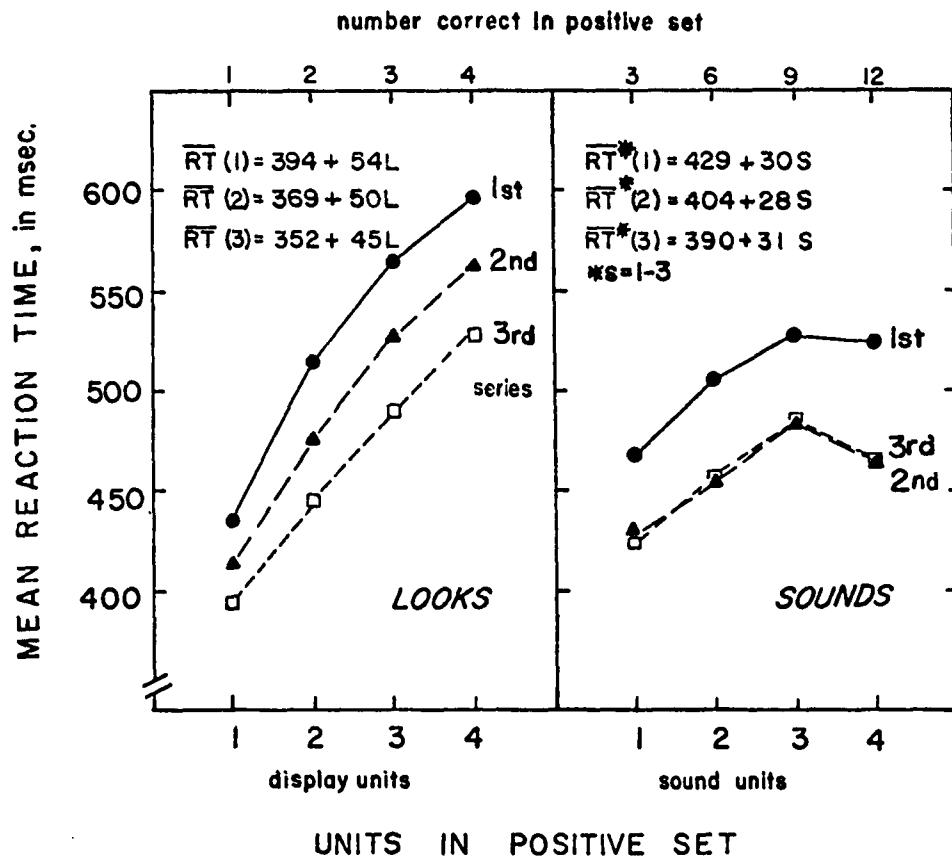


Fig. 2. A comparison of reaction times as a function of units in the Theoretical Positive Set for Looks and Sounds Conditions of Experiment I, displayed by series. Slopes and intercepts for linear regressions, recorded on the figure, were calculated using theoretical units and the equation $y = a + bx$.

and 4, and extrapolated to the one-item case. Moreover, the slope of this new line for Series One is 45 msec./item, which is identical to the slope in Series Three. [Wattenbarger also found the Yes response in the one-item case to be faster than expected -- the discrepancy being approximately 90 msec.] In the present study, the No latency was also faster for the one-item trials in Series One, but only 20 msec. below the expected time. This difference also disappears by Series Three. In conclusion, practice does not appear to affect the slope, which we may take to be about 45 msec. per visual item, but practice does affect the intercept and the one-item response time.

In summary, the faster overall reaction time and the shallower slope of the Sounds Condition, relative to the other two conditions of Experiment I, supports the Auditory Hypothesis.

Criterion II

The second criterion suggests the inter-trial, intra-set variability within the Sounds Condition (defining Positive Sets via the Auditory Hypothesis) should be less than the variability within the other two conditions if memory is solely auditory. Table 2 reports two measures of variability: the intra-set variability or standard deviation of trial means within each value of the Theoretical Positive Set, and the intra-set range or difference in milliseconds between the fastest and slowest trial means. On both measures the Sounds Condition is superior, further supporting the Auditory Hypothesis.

Criterion III

Ideally it would be preferable to reorder the data of each condition according to the other two "less compatible" hypotheses, such that all three hypotheses were compared within each condition. In practice, this operation is difficult. Moreover, the slow reaction times, irregular curve as a function of set size, and extreme variability within the Means Condition when ordered by the Semantic Hypothesis eliminates the Semantic

since the Means Condition trials were divided among four sets under the Semantic Hypothesis, relative to dividing the same number of trials among eight sets with the Visual Hypothesis. In general, the more sets, the fewer cases within each set and the lower the intra-set variability. The extreme case -- having a one-trial set, hence zero variance -- does exist in the present reorganization. Nevertheless, by noting that the newly-formed set size six has the same range as the maximum range under the Semantic Hypothesis (286-msec. difference between fastest and slowest trials), we can conclude that the Visual Hypothesis has not contributed to the predictability of the Means Condition data.

In applying the Visual Hypothesis to the Sounds Condition, it will be instructive to review Figure 2. The abscissa for the Sounds Condition (right side of figure) is plotted in auditory units on the bottom axis, but in visual units (number of potentially correct items) at the top of the figure. Since there are always three versions of each sound, the ratio of visual to auditory units is always 3:1. This eliminates the need to regroup trials. All trials remain in the same sets; we merely count each of the four sets as having three times as many memory codes (i.e., 3, 6, 9, and 12 visual units). Consequently, the slope is effectively 1/3 the value listed on the figure (previously calculated in auditory units), or approximately 10 msec./item. Once again, while this visual slope is very shallow, it does not match the slope of the Looks Condition, the latter being 4.5 times as steep. In summary, the Visual Hypothesis fails to meet Criterion III when applied to both the Means and Sounds Conditions. Therefore, it is safe to conclude that memory is not solely visual.

Two Codes?

Although we can eliminate both the Semantic and Visual Hypotheses from consideration as the sole form of memory codes, it is not yet possible to conclude that subjects fail to use a visual form of memory during the Looks Condition. The evidence for this position comes from the two previously cited findings: (1) the one-item trials for the Looks Condition are faster than the one-sound trials in the Sounds Condition, and (2) the

slope of the Looks Condition is linear. Let us pursue this argument in greater detail.

If memory were solely auditory, then it should require the same time to search a one-sound memory trace regardless of the instructional condition. However, in the Looks Condition when a single item is correct, the subject must first store the sound of that item, and also a code for which version of that sound is correct, since there are three homonym forms for each sound. We shall refer to this additional code as a "tag." It is further assumed that it requires time to check this tag, relative to those instances where any version of the sound could be correct. Consequently, it should take slightly longer to search for a single item in the Looks Condition than a single sound in the Sounds Condition. Figure 3, which displays data from the Looks and Sounds Conditions separately for each subject, reveals that the reaction time for a single item in the Looks Condition was always faster than for one sound in the Sounds Condition. We should note that this finding has an ambiguous interpretation, since it could imply either (1) there is a visual memory which operates in the Looks Condition, although not in the other two conditions, or (2) the intercept of the Sounds Condition is higher than it would be if there were no homonym versions of each sound. The slope and intercept of the Sounds Condition in the absence of homonym stimuli will be evaluated in Experiment II.

ERRORS

Two questions concerning errors deserve our attention: (1) Do the error rates differ among conditions within the same experiment? (2) Does the pattern of confusions favor one or more of the competing hypotheses?

Overall Rates

In answering the first question, we consult Table 3. Although our concern at the present time is only with Experiment I, for convenience the overall errors for all conditions of all experiments are included. In addition to recording total errors, Table 3 reports the proportion of false positive to false negative errors. A false positive is a Yes response to a test item which should have been classified as No -- that is, the item is not a member of the Positive Set. Similarly, a false negative is an incorrect No response to a member of the Positive Set. The ratio of incorrect Yes to No responses is reported in the right-most column of Table 3.

In Experiment I, there is little discrepancy between conditions in total errors. The slightly higher rate for the Looks Condition (3.5% vs. 2.9% and 2.8% for the other two conditions) can be fully attributed to the first session performance of a single subject who made 41 errors. Therefore, we conclude the overall error rates did not produce evidence favoring any of the conditions of Experiment I.

Another important pattern of errors is the change in error rate as a function of the size of the Positive Set. Table 4 reports the proportion of errors according to the size of the Positive Set for all but the Means Condition. As in the findings of Sternberg and Wattenbarger, error rates increase as the number of to-be-memorized items increases.

TABLE 4
PERCENTAGE OF ERRORS BY CONDITION AS A FUNCTION
OF THE SIZE OF THE POSITIVE SET

Condition	Experiment	Units in the Positive Set					
		1	2	3	4	5	6
Looks	I	2.0	3.8	3.5	5.1		
	III	1.3	2.0	2.0	3.3	4.5	10.5
Sounds 1:3	I	2.8	2.9	3.5	2.4		
	II	3.1	4.3	4.3	2.4		
Sounds 1:1	II	2.0	4.1	5.0	5.0	6.6	8.0
	Ext.	3.3	4.5	4.9	5.1	6.7	7.0
Type 1:3	II	3.8	6.9	6.0	3.2		

TABLE 5

OBSERVED NUMBER OF ERRORS AND UNBIASED EXPECTANCIES
WITH REFERENCE TO AUDITORY CONFUSIONS

Error		Looks		Means	
		observed.	expect.	observed.	expect.
(TEST ITEM)					
False Positive	Homonym	142	(77)	146	(77)
	Non-homonym	24	(89)	20	(89)
(STIMULUS SET)					
False Negative	Complete Triplet	7	(16)	11	(10)
	Incomp. Triplet	216	(207)	136	(137)

TABLE 6
OBSERVED NUMBER OF ERRORS AND UNBIASED EXPECTANCIES
WITH REFERENCE TO VISUAL CONFUSIONS

Error	Condition	First Letter of Test		Arabic Number	
		Same	Different		
False Positive	Looks	observed	28	59	55
		expected	(33)	(65)	(44)
	Means	observed	35	85	26
		expected	(33)	(67)	(45)
False Negative	Looks	observed	71	77	75
		expected	(74)	(74)	(74)

only meaningful when both the test and the member of the Positive Set were words. If the test was a word, but the Positive Set homonym was an Arabic number, the test was classified "different." If the test item was an Arabic number (and a homonym), it was listed separately (right-most column). False positives to both words with a different first letter and to Arabic numbers (where the Positive Set homonym was necessarily a word) represent contradictions of the Visual Hypothesis, since neither could be visually similar to items requiring a Yes response.

The Visual Hypothesis was also tested for false negatives in the Looks Condition, as shown in Table 6. If words are encoded according to their visual pattern with special attention to the first letter, we expect an error in either (1) the recognition of a test word beginning with the same first letter as the stimulus (misperception of the test), or (2) the initial encoding of the stimulus, such that another word beginning with the same letter was stored in its place (misperception of the to-be-memorized word). Consequently, the Visual Hypothesis predicts more false negatives to words which have another stimulus involving the same first letter (as TWO-TOO, FOUR-FOR) relative to either unique-first-letter words (as ONE, WON, EIGHT, ATE) or to Arabic numbers. There is no bias among false negatives for any of these three equally-represented categories, further disconfirming the Visual Hypothesis.

Semantic Hypothesis. The Semantic Hypothesis was evaluated using the 55 false positives of the Looks Condition where the test item was an Arabic number. Since the Arabic homonym version should have received a No response, the Positive Set must have displayed (1) the non-number version of the sound, or (2) the spelled number version of the sound, or (3) both the non-number and spelled versions. If memory preserves the meaning of the stimulus, the incorrect response to the Arabic would occur more often when the stimulus set contained the spelled number than when it contained the non-number (since the former is a synonym). Table 7 reports the observed frequencies of each Type of stimulus set. The obtained values of 11, 21, and 23 errors depart significantly from expectancies based on the tally of 20, 13, and 15 trials containing non-numbers, spelled numbers, and both

TABLE 7
FREQUENCY OF INCORRECT YES RESPONSES TO ARABIC HOMONYMS BY TYPE
OF MEMORIZED ITEM IN THE LOOKS CONDITION, EXPERIMENT I

Series	Non-Number	Spelled Number	Both
One	6	8	13
Two	2	7	7
Three	3	6	3
Total	11	21	23

Types, respectively. While these data appear to support the Semantic Hypothesis, a more detailed analysis reveals that 11 of the 21 errors to spelled numbers come from only two trials. Trial 42 displayed the items 2, 4, and EIGHT (false positive response to test item "8") while Trial 58 displayed the items ATE, FOUR, 8, ONE, 2, and FOR (false positive response on this practice trial to the test item "1"). In both cases there are two other items in memory coded as Arabics. There were no matching trials where the critical item (i.e., EIGHT for Trial 42 and ONE for Trial 58) appeared in the non-number version. Thus the bulk of the evidence for semantic confusions is heavily dependent on two trials which produced unusual difficulty in keeping track of which version went with which sound. This same evidence supports the alternative explanation that subjects require additional tags to denote which version of a homonym is correct.

Additional evidence that subjects use tags specifying the Type of homonym in question was provided by false positive responses to Null Set Arabics. Four of the seven errors were on Trial 44, where the Positive Set was composed of the items 1, 2, 4, and 8. The average reaction time for the errors was 377 msec., which was 105 msec. faster than the mean Yes time for this trial (for Series One and Two which contained these errors). Another error was on the similar (practice) Trial 57 (Positive Set = 1, 2, 4, 8, ONE, and TWO; response time = 374 msec.). The other two errors were also on trials containing Arabics (Trial 13 = 1, ONE; Trial 10 = 8). Both involved precipitated responses (times of 382 and 272 msec., respectively). While these errors can be interpreted as evidence for the Visual Hypothesis, the lack of converging evidence, the small number of cases, and the extremely brief response times hardly inspires confidence in this interpretation.

CHAPTER IV

QUESTIONS ARISING FROM EXPERIMENT I

This chapter confronts several technical difficulties in interpreting the data from the principal experiment. Five questions provide the impetus for the design of Experiments II and III. In general, these additional experiments increase our confidence in the previous conclusion that memory is auditory.

The findings of Experiment I, for both response latencies and errors were interpreted to support the Auditory Hypothesis. Nevertheless, some of the findings were difficult to explain if memory were solely auditory. In particular, we ask: (1) Could the data which appears to support the Auditory Hypothesis have an alternative interpretation? (2) In the Sounds Condition, why are the response times to four sounds faster than to three sounds? (3) Would the slope of the Sounds Condition be the same if there were no homonym versions for each sound? (4) Why is the average reaction time to one item in the Looks Condition faster than the time for one sound in the Sounds Condition? (5) Was the steep slope of the Looks Condition (relative to the slope of the Sounds Condition) influenced by the presence of other instructions -- specifically, the emphasis on the auditory properties of the stimuli due to the use of homonyms?

The last three questions are directly resolved by developing a set of stimuli which include no homonyms and then presenting the materials with both the "sounds the same" instruction (Experiment II, Sounds 1:1 Condition) and the "looks the same" instruction (Experiment III). We shall consider these experiments presently.

The first two questions require that we inspect a reasonable alternative to the Auditory Hypothesis. In Experiment I, an analysis of a set of trials suggested such an alternative explanation. Four trials displayed all three versions of a homonym triplet (the trials were numbered 31, 32, 33, and 34, although they were scattered in the trial sequence). For all three instructions the subject's task was the same -- to respond Yes when

any one of the three members of the triplet appeared as a test item. Consequently, we expect the average response latencies for these trials to be the same across the three conditions of Experiment I. The obtained mean pooled response times for the Looks, Means, and Sounds Conditions were 466, 460, and 427 msec., respectively. These differences were largely due to Yes response latencies, which produced times of 455, 441, and 404 msec., compared to No response times of 477, 478, and 451 msec. (same order of conditions in each case). For all six subjects the Yes and No response times were fastest in the Sounds Condition. Note that the rank ordering of conditions parallels the number of "chunks" or "groupings" implied by the instructions. The Sounds Condition implies the three stimulus items having the same sound are chunked into a single unit, so each of the Trials 31-34 would represent one chunk. The Means Condition would take two chunks to specify the three items (numbers and non-numbers); and the Looks Condition would need three chunks. The fewer chunks, the faster the response latency. Since chunking items is known to facilitate recall (Miller, 1956), perhaps a part of the effect we have observed in the Sounds Condition was not due to the auditory properties of the items at all. Maybe it was merely the result of a learned habit of chunking classes of items. We now entertain an alternative model to the Auditory Hypothesis which may account for the same findings. The "Chunking" Model proposes that the greater efficiency for responding with large numbers of potentially correct items in the Sounds Condition, relative to the Looks Condition, is due to the application of a chunking rule, not an auditory memory code.

The Chunking Model does not take a position concerning the specific nature of a memory code with reference to visual, semantic, or auditory properties. It merely assumes that units of memory can be clustered together via a new (abstract?) code. Consequently, we do not have a fourth hypothesis, as such. The possibility of chunking is an independent consideration which can serve as an extension of any of the three major hypotheses. The present concern is that a chunking strategy might have contributed to the flat slope of the Sounds Condition, in effect undermining the rationale for Criterion I. The obtained slopes eliminate both Visual + Chunking and Semantic + Chunking possibilities, but leave three alternatives: (1) purely

Auditory codes, (2) Auditory + Chunking, (3) solely Chunking of abstract codes. Experiment II was designed to differentiate these alternative explanations.

How might chunking be effective? We might assume that subjects can group items into semantic equivalents and into auditory equivalents. The stimuli of Experiment I contain eight "meaning-chunks" or four "sound-chunks." Furthermore, we assume the subject can discriminate those items which are potential stimuli from those of the Null Set. For example, the Chunking Model interprets the fast reaction times on the four-sound trials in the Sounds Condition as the result of the more general two-set discrimination. The Chunking Model proposes that the subject learns to discriminate those items which are possibly correct stimuli from those of the Null Set. The four-sound trials are unique in that any item which is ever correctly classified as Yes will require a Yes response during these trials. Consequently, the subject need only search through one memory chunk on the four-sound trials. Consistent with this interpretation is the finding that the drop for the four-sound trials, relative to the three-sound trials, increases with greater familiarity for the stimulus items (i.e., across series). Figure 3 verifies that all six subjects were faster to four than to three sounds. However, no subject was as fast on four-sound trials as one-sound trials. Since both sets are presumed to involve a single chunk, the slower time to four-sound trials is inconsistent with an extreme version of the Chunking Model.

Another difficulty with the Chunking Model arises in the interpretation of the linear slope of the Sounds Condition for Positive Set sizes 1-3. The argument associated with sets 1-4 goes as follows. Suppose the subject has five chunks or categories: one for each of the four homonym triplets, and a fifth for the Null Set. Then, on those trials with a Theoretical Positive Set of two units, it would be most efficient simply to recall the two correct sounds. However, on the trials where three sounds are memorized, it would be most efficient to store just two chunks -- the homonym triplet which is incorrect, and the Null Set. Thus both types of trials would require memorization of two sets or two chunks. This reasoning leads to an expected function shaped like an inverted "U" -- equally fast times for set sizes 1 and 4, and equally slow times for set sizes 2 and 3. The data of Experiment I indicate that this was apparently not done.

The data from a new condition, designed to promote chunking according to some rule other than auditory similarity, could take one of three forms: (1) identical to the Sounds Condition results, which would invalidate the 'pure' auditory code interpretation, (2) similar to the expected results of the Chunking Model (inverted "U"), but unlike the Sounds Condition, which would establish that chunking can take place, that chunking could have produced the faster latencies on four-sound trials due to a two-set discrimination, and that the previous interpretation of auditory memory codes remains unchallenged by the Chunking Model, and (3) altogether different from the two expected results, which defies interpretation. The Type 1:3 Condition of Experiment II was designed to evaluate these possibilities.

EXPERIMENT II

Conditions

The three conditions of Experiment II included: (1) Sounds 1:3, a replication of the Sounds Condition of Experiment I, (2) Sounds 1:1, a "sounds the same" instruction applied to new stimulus materials devoid of homonyms, and (3) Type 1:3, a rearrangement of the stimulus materials of the original Sounds Condition designed to evaluate the Chunking Model. Each subject was exposed to each condition only once.

The Sounds 1:3 Condition (the 1:3 representing the maximal D:C ratio) was included to serve as a control-baseline for interpreting the other two conditions of this experiment. It differed from the earlier trial sequence in that the counterbalancing equated the number of trials for each value of stimulus variable C, rather than variable D, as used in Experiment I in order to favor comparisons within the Looks Condition. Consequently, there were an equal number of trials representing 1, 2, 3, and 4 sounds. Up to six items were displayed, permitting D-C pairs of 1-3, 2-3, 3-3; 2-6, 4-6, 6-6; 3-9, 4-9, 6-9; 4-12, 5-12, and 6-12.

The Sounds 1:1 Condition used words which did not have homonyms plus the non-numerical items from Experiment I. As a result, there was one item correct for each item displayed (thus the D:C ratio was 1:1). The size of the Positive Set was expanded to six items, yet with 12 potential stimuli, there were always residual members of the set of potential stimuli which could serve as No response items on a given trial. This eliminated the possibility of a two-set discrimination effect. All items were three-letter words, which avoided discrimination based on word-length. [The mean reaction time to Arabic numbers in the one-item Looks Condition in Experiment I was 21 msec. faster than the time for either Type of spelled word.] The set of potential stimuli which contributed to the Positive Set for the Sounds 1:1 Condition included: WON, TOO, FOR, ATE, CAT, DOG, COW, FIG, ARM, LEG, EAR, and TOE. The Null Set items were: BOB, JIM, TOM, JOE, FLY, ANT, BEE, BUG, OAK, ELM, ASH, and PEA.

The Type 1:3 Condition reordered the original stimulus matrix of four sounds each with three Types, into a matrix involving four Types each with three sounds. Three Types were the same as in Experiment I -- Arabic numbers, English spelled numbers, and English non-number words -- to which Roman numerals were added. The pool of potential stimuli included the items: 1, 2, 4; I, II, IV; ONE, TWO, FOUR: WON, TOO, FOR. Since the subject's task was to respond Yes to any Arabic number if an Arabic appeared in the to-be-memorized display, the Null Set could not contain any Arabic numbers. The intention was to make the stimulus and Null Sets as distinct as possible, and to provide labels to the subjects to help define the sets. This should have allowed the Type 1:3 Condition every advantage in being as efficient as the Sounds 1:3 Condition, according to the Chunking Model. If the results show faster reaction times for the Sounds Condition, we will conclude that the Auditory Hypothesis better describes the previous findings. To maximize discriminability, Null Set items included: &, %, \$, +, =, <, MKP, NHM, LFD, LUI, VIN, MOI.

Subjects were given the same instructions for the Sounds 1:3 and 1:1 Conditions. They were told that previous research showed subjects do best if they pronounce aloud the different sounds in the to-be-memorized list, and that they should rehearse the items silently. The possible Positive

Set sounds were described verbally to each subject before the testing commenced. The instructions for Type 1:3 included an illustration informing the subject that there were four sets of items which had to be remembered, each set containing three members. Since the instruction on the display scope for this condition read, "THE SAME SET," the subject had to respond Yes to any item within a memorized set. The labels "arabic," "roman," "english #," and "english words" appeared beside their respective sets. [Appendix F contains the illustrated instruction sheet.] These labels were recommended to the subject, with the further comment that the subject was free to use any labels he chose, "such as A, B, C, or D -- as long as they work." A sample trial was typed out, and the subject studied the sheet for three minutes before beginning the seven practice trials. Subjects were further instructed to call aloud the sets which were present while the to-be-memorized display was on the scope, and then to rehearse the labels. All subjects used the suggested labels except one who referred to the spelled numbers as "American numbers."

Results

There was considerable variability among subjects within two conditions of Experiment II, creating difficulties of interpretation. However, an analysis of sequential effects (Chapter V) revealed that the data were consistent for non-repetition responses, but varied in an unpredictable manner for response repetitions. In addition to being consistent across subjects and across testing sessions, the latencies of the non-repetition responses agreed with expected values for both slopes and intercepts. Consequently, we will not dwell in the present chapter upon puzzling details which will be resolved in the next chapter.

Our first concern is whether the data from the naive subjects are comparable to those of the trained subjects in Experiment I. The Sounds 1:3 Condition provides this test. When the results of Figure 4 are compared with those of Figure 2, we verify that the average reaction time was nearly identical for equivalent practice. The mean reaction time in Experiment I for Series One, disregarding set size, was 495 msec., compared to 499 msec.

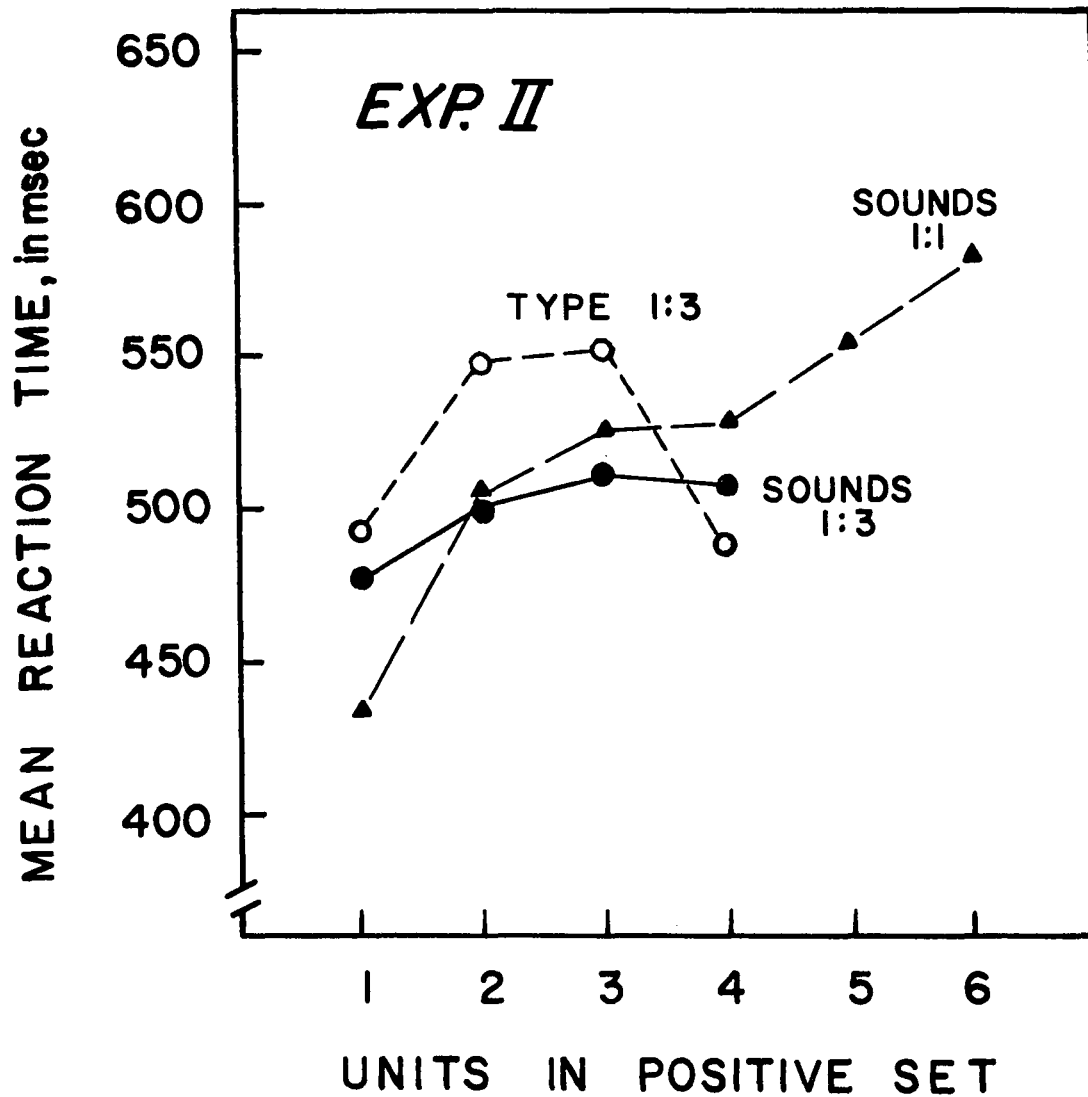


Fig. 4. A comparison of all conditions of Experiment II, plotting response times as a function of the Theoretical Positive Set. The data represent 13 subjects per condition, irrespective of hour.

in Experiment II. The slope in Experiment II across sets 1-3, however, was 17 msec./sound, considerably below the 30-msec. value of Experiment I.

Over comparable trials for both experiments, the Yes response slope in Experiment II (30 msec. per sound) was equivalent to that of Experiment I, although the No response slope (8 msec. per sound) was dissimilar. The fact that the intercepts of Yes vs. No responses differed by 101 msec. in Experiment II (compared to 57 msec. in Experiment I) suggests Experiment II subjects may have concentrated more on making a fast Yes response, such that all No responses tended to be made at more nearly the same latency.

Figure 4 shows the obtained results fit well with the inverted "U" pattern predicted by the Chunking Model for the Type 1:3 Condition. The interpretation of these data is: (1) chunking can affect response latencies in a predictable manner, yet (2) chunking alone can not account for the results of the Sounds 1:3 Condition. While the reaction time to one unit in the Theoretical Positive Set was only 15 msec. slower for Type 1:3 relative to Sounds 1:3, the difference increased to 50 msec. for set sizes 2 and 3. The faster time for the Sounds Condition was obtained for 10 of the 12 subjects serving in both conditions. The two subjects with slower latencies for the Sounds Condition had served in this condition for their first hour. Thus the advantage of the Type Condition for these subjects may simply reflect the effects of practice.

The faster time for No responses with set size four for the Type Condition is consistent with the Chunking Model. We assume the subject's strategy for the four-Type trials involved memorization of the single Null Set chunk, which required the No response, rather than the memorization of the four chunks requiring a Yes response. Although the No response was still slower than the Yes, the difference was reduced from 60 msec. for set size one (memorize one Yes response chunk) to 13 msec. for set size four (memorize one No response chunk). Only the Yes-No difference for set size four failed to reach statistical significance. Also, the No response time for set size four was significantly faster than the No times to all three other sets. By comparison, in the Sounds 1:3 Condition the fastest No response times occurred for set size one.

In addition to faster response latencies, the Sounds 1:3 Condition was superior in accuracy. Incorrect responses were made to 3.5% of the tests for the Sounds Condition compared to 5.0% for the Type Condition.

The results of the Type 1:3 Condition relate to the first two questions posed at the outset of this chapter. First, could the results of the Sounds 1:3 Conditions of Experiments I and II merely reflect the effects of a chunking strategy? The pattern of results from the Type 1:3 Condition led to the conclusion that that condition reflected "pure" chunking. Since the Sounds 1:3 Condition was significantly different from the Type Condition, it must have involved some other kind of processing. We reject the possibility that the results of the Sounds Condition can be solely attributed to chunking. However, we should not overlook the similarity between the Sounds 1:3 and Type 1:3 Conditions. Specifically, both produced a drop in latency for sets containing four units, although this drop was significantly greater in the Type Condition. This similarity suggests that while the Sounds Condition did not solely reflect chunking, it may have been susceptible to a two-set discrimination effect due to the selection of stimulus materials which incorporated a fixed set of four sounds. Therefore, in answer to our second question concerning the drop in latency for the four-sound trials, we conclude this result is an artifact of the stimulus materials. Verification of this conclusion is provided by the absence of such an effect in the Sounds 1:1 Condition which did not permit such a two-set discrimination.

The results of the Sounds 1:1 Condition were somewhat irregular. Under the Auditory Hypothesis, we expect the slope of this condition to be identical to that of the Sounds 1:3 Condition (i.e., 17 msec./sound). Although a slope of 19 msec. per sound was obtained across set sizes 2-6 (26 msec. for sets 1-6), the data points do not adequately match those of the Sounds 1:3 Condition. Since there were fewer visual patterns per sound, the stimulus uncertainty was reduced (i.e., the probability that a given item from the Positive Set would appear as a test item increased). This effect could result in a shorter time interval needed by a recognition stage of processing, in which the test item is transformed into an appropriate code for the memory comparison process. As a result the total time for a response would be reduced, producing faster latencies in the Sounds 1:1

Condition. Yet Figure 4 shows that only set size one resulted in a faster time for the 1:1 Condition, and that point was discontinuous with the rest of the function. In fact, this point was 2 msec. faster than the time for the Looks Condition, Series One of Experiment I. This similarity to the Looks Condition suggests the subjects may have been functioning in the same way during both conditions on those trials displaying a single item.

Large between-hour differences are shown in Figure 5. The term "hours" refers to the order of participating in each of the three conditions of Experiment II. Unlike the between-series comparisons of Experiment I, between-hour comparisons are necessarily between-subject comparisons, since each subject served in each condition only once. For all three hours, those trials with a single to-be-memorized sound produced exceptionally fast reaction times. Moreover, the data from Hours Two and Three with Sounds 1:1 were consistent and parallel to the Sounds 1:3 Condition for the critical set sizes of two and three sounds. By contrast, the data for Hour One appears to be more like the results of the Looks Condition, although the obtained slope is less than 31 msec./sound (24 msec./sound for sets 2-6). One interpretation of these changes relies on the fact that all subjects serving Hour Three in the 1:1 Condition have already served in the 1:3 Condition. Presumably, the latter condition forced a greater reliance on the sound properties of the memorized items. This task may have aided these subjects in the use of auditory representations, relative to those subjects just beginning the experiment. In preview of Chapter V, we can anticipate that these between-hour differences are mainly varied reactions to response repetitions, with the more experienced subjects making faster responses to repetitions than the less experienced subjects.

Before relating the results of the Sounds 1:1 Condition to our five questions, it will be to our advantage to satisfy ourselves that these findings are reliable. The Extension 1:1 Condition fulfills this need.

Extension 1:1

In order to obtain a more stable set of data for the Sounds 1:1 Condition, three subjects from Experiment I each participated for four hours in the Sounds 1:1 Condition. These subjects received the same in-

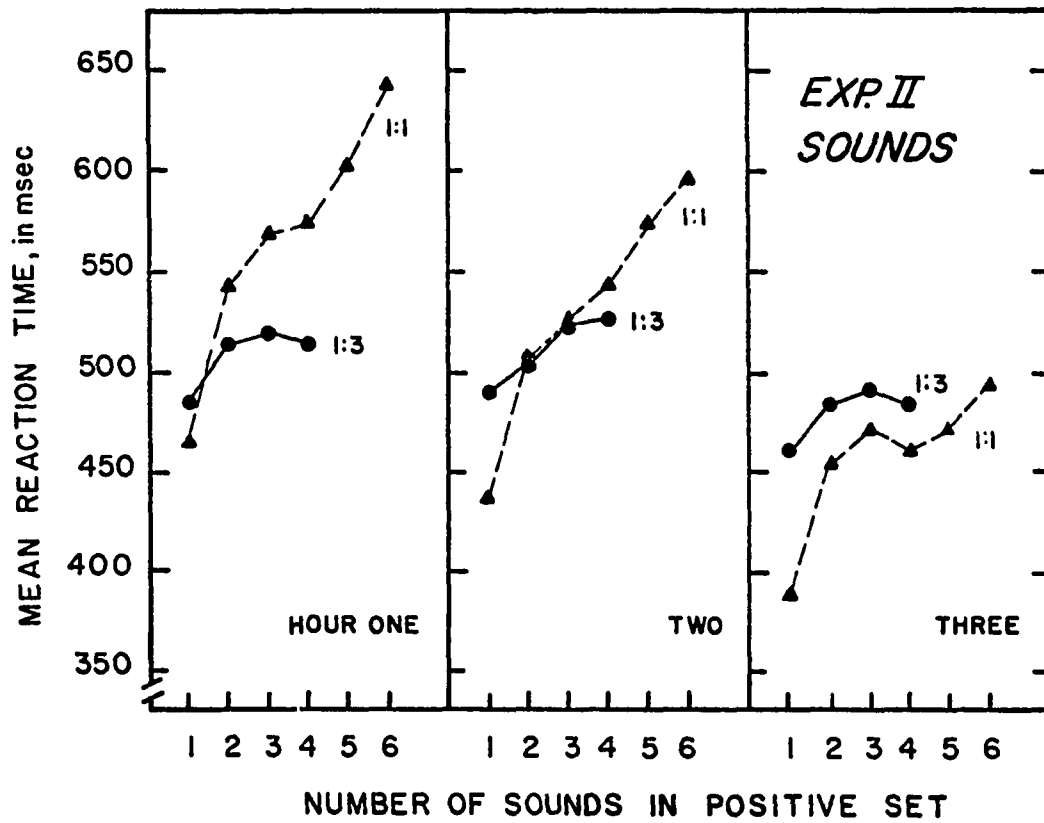


Fig. 5. Between-subject comparisons of the Sounds 1:3 and Sounds 1:1 Conditions of Experiment II, separated by hour.

structions as the subjects of Experiment II, but were encouraged to work faster than they had previously done. [The overall error rate of the Sounds 1:1 Condition was 5.2%, considerably higher than that of the experienced subjects working on the Sounds 1:3 Condition in Experiment I. The error rate of the Extension 1:1 subjects, responding faster than before, was also 5.2%.]

The data obtained from these subjects were extremely stable, showing no effects of practice on either slopes or intercepts. The mean reaction times across all subjects for all four sessions provide the clearest estimate of performance. The mean latency, irrespective of set size, was 77 msec. faster than that of the Sounds 1:1 subjects of Experiment II (439 vs. 516 msec.). However, the slope for set sizes 2-6 was identical to the Sounds 1:1 results -- 19 msec. per sound, accounting for 99.0% of the total variance (slope for sets 1-6 was 25 msec./sound, predicting 93.3% of the variance). The obtained latency for the one-item Positive Set was 39 msec. faster than the time predicted from the linear regression through sets 2-6. The comparable discrepancy for one-item trials in the Sounds 1:1 Condition of Experiment II was 50 msec.

The data from the Sounds 1:1 and Extension 1:1 Conditions relates to questions 2, 3, and 4, as posed at the beginning of this chapter. Question 2 asks why the response times to four sounds were faster than the times for three sounds. The absence of this drop for the Sounds 1:1 Conditions, in conjunction with the chunking effects observed in the Type 1:3 Condition, verifies the conclusion that the four-sound drop was an artifact of the stimulus materials, resulting from the two-set discrimination between stimuli and Null Set items.

Question 3 asks whether the slope of the Sounds Condition would be the same for stimulus materials lacking homonyms. The answer is simply yes, the slope was unchanged. We are not able to conclude whether or not the intercept is affected by homonyms, due to the equivocal findings in Figure 5.

Question 4 asks, "Why is the average reaction time to one item in the Looks Condition faster than the time for one sound in the Sounds Condition?" Although we are still in doubt as to the cause for one-item trials producing latencies which are discontinuous (faster) with the other latencies,

we recognize that the fast reaction time is not restricted to the Looks Condition. Subjects may not need to use a memory comparison stage of processing to make a judgment involving a single target item.

Having dealt with the first four questions, we are now prepared to consider question 5 -- was the slope of the Looks Condition affected by the presence of homonyms? Experiment III was designed to answer this question.

EXPERIMENT III

Experiment III was designed to provide a between-subjects comparison of Looks and Sounds instructions for materials without homonyms. Therefore it was run concurrently with Experiment II, using five new subjects who were selected in the same manner as those in the second experiment. The trial tape was identical to the Sounds 1:1 Condition of Experiment II, except the instruction was changed to read, "LOOKS THE SAME." Subjects were given the further instructions to try not to pronounce the to-be-memorized items as they appeared, but to concentrate on their form and letter configuration. Each subject served for only one hour, after which he was questioned concerning rehearsal. All five subjects reported being unable to prevent silent naming of the items, and most subjects felt they were rehearsing at least part of the time, having some uncertainty about what they had been doing.

The results of Experiment III are displayed on the left side of Figure 6. Plotted with these results are the data from Hour One of the Sounds 1:1 Condition of Experiment II and Series One of the Looks Condition of Experiment I. The two Looks Conditions are nearly parallel except for the dip in Experiment III with set size four. The slope for Experiment III was 42 msec./item, only slightly below the 45-msec. value which best represented the data of Experiment I. The minor difference in intercept can be attributed to the fact that subjects in Experiment III were serving their first hour, while the data from Experiment I has collapsed latencies from subjects in their second, third, or fourth hour of responding. The results of the Sounds 1:1 Condition of Experiment II were quite different from the Looks Condition of Experiment I, the best-fitting slope for the Sounds Condition being less than 24 msec./sound, calculated across the sets of 2-6 items

(accounting for 94.1% of the variance). A further demonstration of the departure of the Sounds vs. Looks data is apparent from the right side of the figure, which illustrates the effects of practice. The steeper slope represents the data from Series Three of Experiment I, incorporating performance after seven, eight, or nine sessions (the variation being due to the different sequences for serving in the three conditions of Experiment I). The shallower slope displays data from the Sounds 1:1 Condition for those subjects working their first hour at this condition, but after two sessions at other conditions. The slope for points 2-6 was only 8 msec./sound. Once again we note the one-item reaction time for the Sounds 1:1 Condition appears to be discontinuous with the rest of the Sounds function, but the same as the one-item response time of the Looks Condition. Despite this minor similarity between the Sounds 1:1 and Looks Conditions, Figure 6 establishes that the Sounds 1:1 Condition differs from the Looks Conditions of both experiments.

In answer to question 5, the small differences between the Looks Conditions of Experiments I and III suggests that the former was relatively uninfluenced by either the presence of homonyms or the other conditions of that experiment.

While several puzzles have been resolved, two remain. (1) Why were the slopes of the Sounds Conditions of Experiment II flatter than those of Experiment I? (2) If memory is solely auditory, why did the Looks instruction of Experiment III produce a steeper slope for the same materials as used in the Sounds 1:1 Condition? This last question opens again the possibility of two forms of memory. In summary, Experiments II and III have resolved several difficulties in the interpretation of Experiment I. In the process, the above two questions arose. Chapter V will resolve these difficulties so that we may return to our original question about the form of memory codes.

CHAPTER V

RATIONALE FOR SEQUENTIAL ANALYSIS

Our primary interest to this point has been to apply the Sternberg paradigm to the problem of identifying the nature of those memory traces operating in short-term recognition memory. Now we direct our attention to the method which has served as a tool for discovery. We ask, "Are the assumptions of a serial exhaustive search justified?" Specifically, we shall test one implication of the serial exhaustive search model -- that the slope should be unaffected by sequential patterns of responding.

The rationale for this prediction was set forth in Chapter I. Three categories of sequential tests were identified: (1) Non-Repetitions (NR), referring to a change from either a Yes to a No response or vice versa, (2) Equivalence repetitions (EQ), where a response is the same as on the previous trial, but the stimulus has changed, and (3) Identity repetitions (ID), where both the stimulus and the response are the same. Differences in intercepts for the three types of sequential responses will not be damaging to the search model, since the Additive Factors Method suggests these effects are produced by other stages of processing. However, differences in slopes among these three response categories would imply the sequential properties of responding affect the memory comparison stage. Consequently, we will concentrate on differences in the slope parameter.

Since the probability of a Yes response was constant ($p = .50$), the relative distribution of IDs and EQs must necessarily vary with the size of the Empirical Positive Set. With one item, all Yes repetitions are IDs. With two items half will be IDs, half EQs. The extreme case of 12 distinct items in the Empirical Positive Set (or four sounds) leads to an overall probability for an ID repetition of $(1/2)(1/12) = .042$ (therefore the probability of an EQ would be the complement, or .458). Increasing the size of the Positive Set will increase the likelihood of an EQ Yes repetition, but decrease the likelihood of an EQ No repetition, due to fewer Exclusive items in the No response pool. The opposite relationship holds for ID repetitions.

Table 8 lists the slopes for the Looks and Sounds Conditions of all experiments, divided into Yes vs. No responses for all three sequential patterns. Only the best-fitting straight lines are recorded. The range for this calculation and the regression correlation of each slope are also noted. The criterion of selecting lines having the highest correlation coefficient serves as a preliminary organizing principle. All cited values will employ this criterion unless noted otherwise. However, we will amend this criterion later in an attempt to establish the most consistent time estimation across experiments for similar instructions. [Appendix G lists the mean reaction times as a function of set size and sequential categories for those conditions listed in Table 8.]

Typical patterns of results can be seen in Figures 7 and 8, both displaying Yes responses, by sessions. Figure 7 shows the Looks Condition of Experiment I while Figure 8 displays the Sounds 1:1 Condition of Experiment II. First we should note that IDs usually produce the flattest slopes, EQs the steepest slopes. The NRs, which constitute half the total responses, provide the most stable slopes. For example, in Experiment I the Looks Yes NRs had slopes of 45, 45, and 49 msec. per item, for Series One, Two, and Three, respectively. The No NRs for these same series resulted in slopes of 41, 45, and 47 msec. per item. In contrast, Yes ID slopes ranged from 18 to 38 msec. for these series and No IDs ranged from 47 to 105 msec. per item -- the latter values being the only exceptions to our generalization that ID responses produce the flattest slopes.

The importance of the stability of the NR slopes is most apparent in Figure 8. The steep slope for Hour One subjects with all responses combined was previously perplexing (see Figure 5). Now it is apparent that much of that effect was contributed by EQ responses. The Yes NR slopes vary within 10 msec. The No response NRs assume best-fit values of 21, 27, and 3 msec. per sound for the three hours -- only the last value being difficult to explain. Therefore two sequential components contributed to the previously noted shallow slope of Hour Three subjects -- a shallow NR slope and a curvilinear ID function which involved shorter latencies with larger Positive Sets.

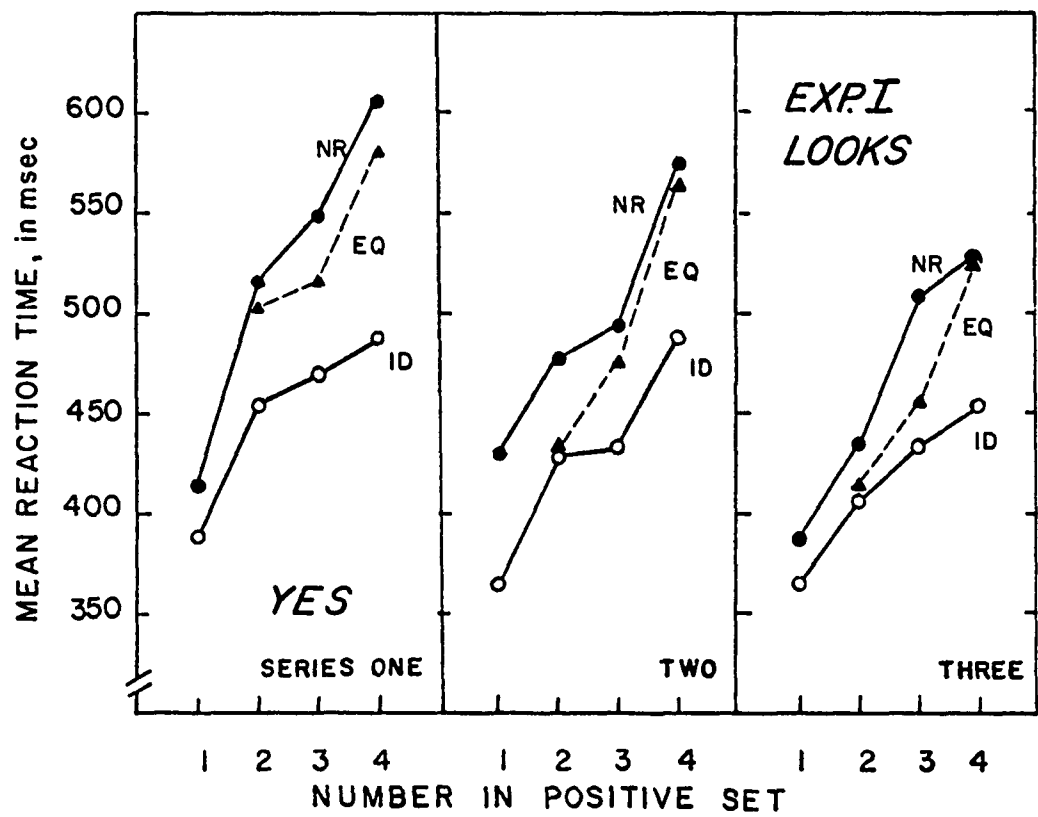


Fig. 7. Sequential effects for Yes responses in the Looks Condition of Experiment I, displayed by series. NR represents Non-repetitions, EQ refers to Equivalence repetitions, and ID refers to Identity repetitions.

The lack of agreement among the three sequential response categories using the slope parameter to characterize the results of any condition can be neither overlooked nor simply explained from the present findings. Such sequential effects are clearly contrary to the predictions from the serial exhaustive search model. The variables of the present study do not appear to provide a clue for understanding these effects, since different patterns emerged across those conditions which have otherwise seemed comparable. At best, the present study has merely called attention to a phenomenon which requires explanation.

IN SEARCH OF A SINGLE COMPARISON TIME

The stability of NR slopes suggests that this component of the reaction time data be used to estimate the average comparison time. The comparison time is thought to represent the time to match the encoded version of the test item with a single memory trace. Without partitioning of sequential effects, the overall slopes in the Sounds Conditions ranged from a high of 30 msec. per sound in Experiment I to a low of 17 msec. per sound in Experiment II (Sounds 1:3). Do the NR slopes more nearly converge on a single comparison time across experiments?

Examining the NR slopes for both Sounds 1:3 and Sounds 1:1 Conditions in Table 8, the time of about 20 msec. emerges as a good estimate for Yes responses, being slightly underestimated for the Sounds 1:1 Condition, of Experiment II, but compensatingly overestimated by the Extension 1:1 data. This time estimate is also a good fit for No responses in Experiment I and the two Sounds 1:1 Conditions. Can these estimates be improved, reducing their range? First we should recall that the single criterion for listing those slopes which appear in Table 8 was the magnitude of correlation. Let us now consider other slope calculations computed over alternative ranges.

The Sounds 1:1 results are most susceptible to modification. Calculating slopes across the full range of memory sets, the Experiment II slopes become 24.5 and 23.2 msec. for Yes and No responses, respectively. The predictability has not fallen drastically by this alteration ($r = .925$ and $.944$, respectively). The Yes slope for Extension 1:1 is an acceptable

24.7 msec. per sound. Once again resorting to a computation across all six sets, the No slope becomes 23.3 msec. ($r = .958$). The agreement among slopes in these two Sounds 1:1 Conditions continues to be remarkable.

Disregarding the 7 msec. No response slope in the Sounds 1:3 Condition of Experiment II (which had given us trouble in Chapter IV), the remaining seven estimates of the auditory comparison time have a mean of 22.8 msec. and a high-low range of less than five msec. The aberrant 7-msec. slope remains a mystery, perhaps related to the extremely slow reaction times for No responses during that condition.

A detailed analysis of sequential effects has verified that the Sounds Conditions, which we expected to be alike, but which differed between Experiments I and II prior to a partitioning of sequential categories, were in fact the same for NR latencies. Thus we have resolved the first puzzle resulting from Experiment II (as stated at the close of Chapter IV). The second difficulty arising from the additional experimentation involved the discrepancy in slopes comparing the Sounds 1:1 Condition of Experiment II with the Looks Condition of Experiment III. Both conditions used the same materials, but the slope of the Looks Condition was greater. Again, looking only at the NR slopes, this discrepancy evaporates. While the slope of the Looks Condition of Experiment I remains unchanged by the sequential partitioning, the slope of the Looks Condition of Experiment III was reduced to a value consistent with those obtained in the Sounds Conditions (i.e., about 23 msec.). Specifically, the slope for combined Yes and No responses was 23.9 msec. per item (Yes slope = 27.3 msec., $r = .982$; No slope = 20.4 msec., $r = .907$). These calculations exclude those trials displaying either a single item or six items. [The exclusion of set size six is reasonable for this condition since these trials produced twice as many errors as those with five items -- 4.5% vs. 10.5%.] By contrast, the flattest NR slopes for the Looks Condition of Experiment I were 45 (Series Two) and 41 msec. (Series One) for Yes and No responses, respectively. Therefore, we conclude the results of the Looks Condition of Experiment III reflect the same type of processing used in the Sounds Conditions. By restricting our attention to NR slopes, the direct comparison of different conditions and of separate experiments suggests we are measuring the same process -- a serial, exhaustive search through auditory memory codes.

CHAPTER VI

CRITERION III FOR THE AUDITORY HYPOTHESIS

We return to our initial strategy of determining which hypothesis best accounts for the results of Experiment I. The analysis of errors and the supplementary evidence discussed in Chapters IV and V have served to bolster our confidence in the Auditory Hypothesis. We have already noted the fast absolute reaction times, the shallow slope, and the low intra-set variability of the Sounds Condition, providing evidence for the Auditory Hypothesis (consistent with Criteria I and II). Criterion III requires that the Auditory Hypothesis reorder the trials of the Looks and Means Conditions into Auditory Positive Sets such that the newly calculated reaction time slope for each condition be the same as that obtained for the Sounds Condition. Additionally, the resulting intra-set variability should be less than those values reported in Table 2, where these conditions were ordered according to alternative hypotheses. Encouragement to apply Criterion III to the Auditory Hypothesis derives from the quantitative similarity of auditory confusions comparing the Looks and Means Conditions (Table 5). Since we observed nearly identical numbers of false positive homonym errors for these two conditions, we have reason to believe that the memory code functioned the same way in each and every condition of Experiment I.

The strategy for reanalysis of the Looks and Means Conditions is similar to the initial proposal for ordering trials, as outlined briefly in Chapter III. Once again, the weighted response times for all six subjects were combined for each trial. All trials which have the same number of sound units are grouped together. Since these groupings are based on a theoretical interpretation of equivalence (as opposed to using the objective stimulus variables of C or D), it is imperative that these classification rules be clearly stated.

The classification was based on a modified form of the Auditory Hypothesis. It may be termed the Sound-Tag Model. The model proposes that memory traces are isomorphic with the sound of a spoken word. When the stimuli have unique sounds (no homonyms), we conceive that subjects make a serial exhaustive search through a memorized list of sounds. When homonyms are introduced and the different versions of a sound belong to separate categories, it is not sufficient to store the sound alone. We assume that the subject stores the sound, but in addition, he also stores a "tag." A tag can be conceived as a private code which tells the subject which version of the homonym belongs in the Empirical Positive Set. Just as it takes time to compare a test to a memorized sound, we assume it takes time to read the tag. The comparison time for a tagged sound will be longer than the time to check a sound without a tag. The comparison time for the tag can be calculated by simply subtracting the slope across trials requiring tags from the slope across those trials where the sound alone is sufficient for correct classification. If there is no slope difference between these sets of trials, we have no evidence that a tag was used.

We further need to consider two possibilities. Suppose that one trial contains the set "1, 2, 4" whereas a second trial contains the set "ONE, 2, FOR." The instruction is "looks the same." Both sets contain three sounds. But the first set involves only one Type -- Arabic numbers -- whereas the second set involves a mixture of Types. That a subject's performance reflects a recognition of the difference between these sets was suggested by the error analysis. False positives to incorrect Arabics occurred more often on those trials which displayed all Arabics. The Sound-Tag Model separates the two sets by proposing that trials involving homonyms which have a common feature (as when all items are Arabics) may require a single tag for the entire set, rather than a separate tag for each sound. This type of trial will be termed a "one-tag" trial, since we will assume the subject needs to store only a single tag regardless of the number of memorized sounds. When there is a mixture of Types, we assume that the subject must attach one tag to each and every sound. These trials will be termed "one tag/sound" trials.

Ideally, the one tag/sound category would include only those trials where there are an equal number of Types and sounds. This would insure that each sound had a unique tag. Otherwise, subjects might cluster sounds and tags. For example, the Positive Set of 1, 2, FOR, ATE might be organized hierarcnically with two binary classifications: first a search through the two tags (Arabic vs. non-numbers), then a search through the two sounds using the selected tag. Williams (1971) has demonstrated that subjects can use this hierarchical strategy in the memory search task. Since the present study used up to four sounds, but only three different versions of a sound, the ideal exemplars for the one tag/sound category were not available. In general, the one tag/sound category will include fewer Types than sounds for the larger Positive Sets.

The definition of a tag differs for the two instructions in Experiment I. For the Looks Condition a tag is needed for each of the three Types, while in the Means Condition a tag denotes one of the two semantic forms of a sound. Trial 25, displaying the items ONE, 2, illustrates this point. For the Looks Condition, this trial would be categorized as a one tag/sound trial. However, in the Means Condition, it would be categorized as a one-tag trial since both sounds use the same code (i.e., number versions). Table 9 lists those trials assigned to each category for the present analysis.

The predicted slope for the one-tag trials differs from that of the one tag/sound trials. If a single tag is sufficient, a single time increment is expected, since the search through additional sounds would take the same comparison time as in the non-homonym (therefore no tag) situation. The predicted slope is parallel to the non-homonym slope, but has a constant increment associated with the tag. We expect the slope of the one-tag trials to be about 23 msec. per sound -- the average value for the various Sounds Conditions as calculated in Chapter V. However, the intercept for the one-tag trials should be higher, reflecting the time to examine one tag.

The one tag/sound trials should have a steeper slope than the one-tag trials, since each comparison time now includes the tag-check time added to the sound-match time. We assume the tag would be checked regardless of the outcome of the sound match, consistent with the concept of an

TABLE 9
 LIST OF TRIALS FROM THE LOOKS AND MEANS CONDITIONS
 OF EXPERIMENT I WITHIN EACH SOUND-TAG CATEGORY

Theoretical Positive Set	Looks Condition	Means Condition		
One Item	1,2,3,4,5,6,7, 8,9,10,11,12	3,6,9,12		
	<u>Complete Triplet</u>	<u>Complete Triplet</u>		
One Sound	31,32,33,34	31,32,33,34 18,21,24*		
Two Sounds	--	50,54,55*		
	<u>One-Tag</u>	<u>Tag/Sound</u>	<u>One-Tag</u>	<u>Tag/Sound</u>
Two Sounds	26,29,30	25,27,28	25,26,28 29,30	27,40
Three Sounds	36,38	42,43	36,38,42, 49	43,48
Four Sounds	44,45	46,47	44,45,46	47

* All three versions of the triplet were inferred to be correct, although not all versions were displayed. Trials 31-34 displayed all three forms of the homonym triplet.

exhaustive search. If we assume the time to check a tag is equal to the time to match a sound, the slope of the one tag/sound trials would be twice that of the one-tag trials. Wattenbarger's finding of nearly a 2:1 slope differential for Physical Identity trials (presumably requiring a tag for upper vs. lower case) relative to the Name Identity trials supports this "ball-park" estimate.

There are two ways we can estimate the time required for a tag check. The first is to subtract the intercept of the sounds-alone (no tags) trials from the intercept of the one-tag trials. The other method subtracts the slope of the one-tag trials (or sounds-alone trials, since the slopes should be identical) from the one tag/sound trials, since the slope measures the time for a one-item comparison (sound comparison vs. sound + tag comparison).

In partial summary, the Sound-Tag Model predicts: (1) The slopes for the one-tag and one tag/sound trials will be linear. (2) One-tag trials should have the same slope as the Sounds Conditions, namely about 23 msec. per sound. (3) The one tag/sound trials should produce a steeper slope than the one-tag trials. (4) The lines defining each function should converge at a Theoretical Positive Set size of one sound unit, since at this point each function represents the time to search one sound plus one tag.

The overall mean trial times were first partitioned among the two sound-tag categories, and then grouped according to the size of the Theoretical Positive Set. Figure 9 displays the results for both Looks and Means Conditions. Let us first consider the data from the Looks Condition (left side of figure). With one sound in the Theoretical Positive Set, trials were divided into those displaying a single item (open square) vs. those displaying all versions of a homonym triplet (open circle). They differ by approximately 50 msec., indicating once again that the one-item trials may involve different processing. Part of this difference may be due to marked differences in the probability of a particular Positive Set stimulus ($p = .50$ for one item vs. $.17$ for homonym triplets). Theios, Smith, Haviland, and Traupmann (1971) have shown that stimulus probability plays a major role in determining the reaction time during a memory search task. In addition to the contribution of stimulus probability, the distribution of sequential response categories differed (all Yes response repetitions were IDs for one-item trials vs. 1/3 IDs, 2/3 EQs for one-sound trials).

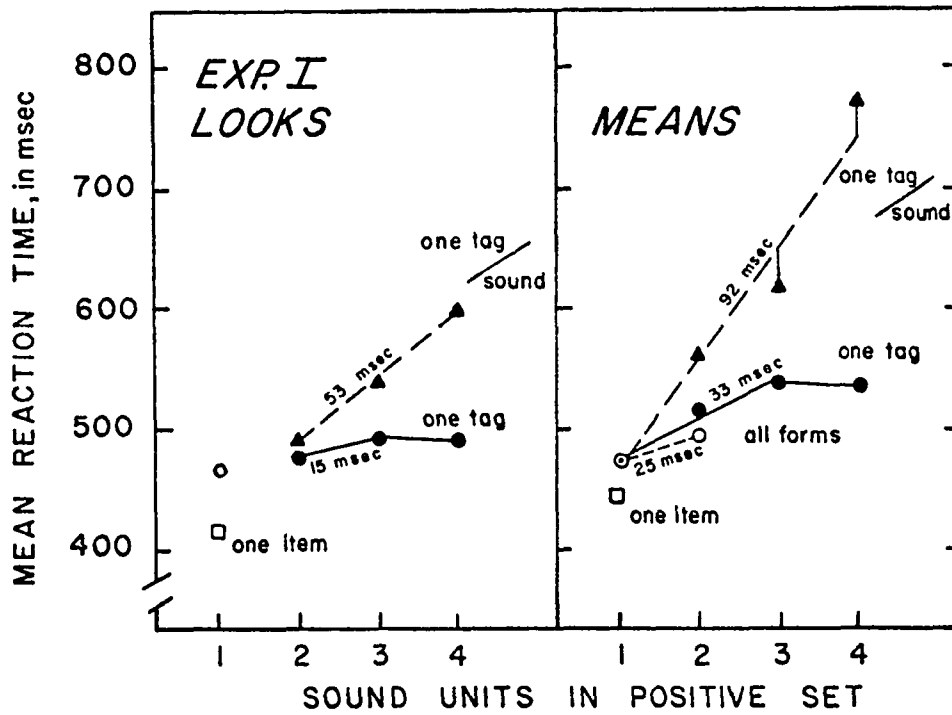


Fig. 9. Reorganization of trials from the Looks and Means Conditions of Experiment I according to the Sound-Tag Model, displayed in terms of the size of the Theoretical Positive Set and separated according to the presumed number of tags needed to specify the correct stimuli.

The one-tag trials result in a function similar to that of the Sounds Condition, including the drop with all four sounds. The increment from set size two to three was 15 msec., somewhat below the slope from the Sounds Condition. [This value was derived from weighted means, rather than the unweighted values in Appendix B.] This curve is markedly different from the one tag/sound function which has a slope of 53 msec. per sound. The two functions converge near the one-sound location, as predicted.

Data from the Means Condition were similar. The one-item trials were faster than the one-sound trials. However, the comparison comparable to that with the Looks Condition produced only a 16-msec. difference between these two types of trials, whereas this difference had been 50 msec. in the Looks Condition.

With the Means instruction, there were a set of trials where all three versions of a homonym triplet could be correct for both one and two sound units in the Positive Set (labelled "all forms" on Figure 9). The difference between the one- and two-sound trials was 25 msec., very close to the expected comparison time of 23 msec./sound. This calculation compares those trials where all forms were inferred to be correct (e.g., displaying 1, WON implies 1, ONE, WON), although some were not displayed. Displaying all members of the triplet (e.g., 1, ONE, WON) in the one-sound case resulted in a 9-msec. faster response time relative to the inferred one-sound trials.

The one-tag curve also approximates the slope from the Sounds Condition. The slope of 33 msec. per sound is somewhat steep, but the presence of greater numbers of faster ID responses for the one-unit trials may have biased this estimate. The difference between set size two vs. three (less susceptible to repetition imbalance) was, encouragingly, 24 msec.

The one tag/sound trials produced the very steep slope (92 msec. per sound) on the right side of Figure 9. The steepness was contributed largely by the point representing four sound units in the Theoretical Positive Set, which, unfortunately, was determined by a single trial (Trial 47). A more reasonable estimate of the one sound/tag function can be obtained using sets two and three, which produce a slope of 57 msec. per sound, similar to the

one tag/sound slope for the Looks Condition (54 msec.). As predicted by the Sound-Tag Model, the line through these two points intersects the one-tag function, when determined by sets of two and three sounds only, at the location of set size one. This intersection point, which is higher than the plotted one-sound point, is 45 msec. slower than the one-item trials. The 45-msec. value more nearly agrees with both the one-sound vs. one-item difference of the Looks Condition (50 msec.), and the obtained vs. predicted discrepancies for the one-item trials in the Sounds 1:1 (50 msec.) and Extension 1:1 Conditions (39 msec.), as reported on page 65.

As noted earlier in this chapter, the tag-check time can be estimated in two ways: by comparing intercepts for the one-tag vs. no-tag trials, and by comparing slopes for the one-tag vs. one tag/sound trials. Since the latter comparison can be made for both the Looks and Means Conditions, let us complete these calculations first. In the Looks Condition, the slope differential was $53 - 15 = 38$ msec. This value may overestimate the tag-check time due to the underestimation of the one-tag slope (expected to be 23 msec., which would yield a value of 30 msec. for the tag-check time). For the Means Condition the most reasonable estimate compares the slopes across only sets of two and three sounds. This difference is $57 - 24 = 33$ msec. per tag.

The comparison of intercepts can be completed only within the Means Condition. The difficulty with the Looks Condition is that we can not calculate a slope (and hence derive an intercept) for the no-tag trials, since there is only one set size where all three versions of a homonym triplet are correct. Even in the Means Condition the no-tag slope can only be estimated by piecing together two components of the difference of interest. We wish to determine the difference in reaction time between trials requiring one tag and trials where all members of the homonym triplet are displayed. We derive this value by getting the difference between one-tag and no-tag slopes where the latter trials have all homonym versions inferred. To this we add the difference in latency between trials where all versions are inferred vs. all versions displayed. Since the slopes of the one-tag (through sets 2 and 3) and of the no-tag or all-forms trials (through sets 1 and 2) are virtually identical, we can measure the difference between the two functions at set size 2. This difference is 19 msec.

The difference between the function where all forms of a sound are inferred vs. where all forms are displayed is 9 msec. Therefore, the difference between the one-tag and no-tag trials is estimated to be $19 + 9 = 28$ msec. We assume this estimate reflects the time to search one tag. Considering this 28-msec. estimate along with the above estimates, we arrive at a tag-check time of about 30 msec., compared to about 23 msec. to search a single sound representation. These estimates do not depart sufficiently to reject the conclusion that the tag-check time is equal to the sound-match time.

In conclusion, the Sound-Tag Model has effectively partitioned the data into two functions: one conforming to the function obtained in the Sounds Condition -- thereby satisfying Criterion III -- and the second derived from the former function by adding the time for tag checks. It was this one tag/sound component which contributed heavily to the steep slopes of the Looks and Means data in the results of Experiment I. We now have good reason to believe that subjects use an auditory code regardless of the instructions. Thus our Sound-Tag Model has synthesized the apparent discrepancy among conditions in Experiment I.

The intra-set variability, calculated for each function separately, must necessarily be reduced in terms of the components of analysis. The actual calculation of variability is not practical due to the small number of exemplars at each set size for each category (see Table 9). Nevertheless, it is intuitively clear that the variance has been partitioned into two segments, reducing the range of reaction times within each segment. In summary, the Sound-Tag Model allows us to reasonably satisfy Criterion III, leading to the conclusion that short-term recognition memory relies on an auditory form of memory code.

CHAPTER VII

CONCLUSIONS

We have probed short-term recognition memory with a reaction time procedure which claims to reveal some temporal aspects of memory-related central processes. Our metaphorical excursion through three experiments began with three conjectures about the properties of memory. We entertained the possibilities that different forms of memory codes function during a word recognition task. The three forms were assumed to preserve the auditory, visual, or semantic features of the Positive Set items. The evidence strongly points to the conclusion that short-term recognition memory codes are basically auditory for verbal materials. Now the time has come to review this evidence and to examine its soundness.

Six facts unequivocally support the conclusion that memory codes are auditory in the Sternberg stimulus classification task:

1. The fastest reaction times occurred in the Sounds Condition of Experiment I. This result is even more impressive given that this condition required the memorization of the largest number of physically distinct visual patterns. The probability that any one of the Positive Set items could occur as a test item was lowest in this condition. This latter finding stands in direct contradiction to the claim of Theios et al. (1971) that the stimulus probability is the critical variable determining reaction times in the memory search paradigm. Although the Sounds Condition was at the greatest disadvantage in terms of both the number of to-be-memorized stimuli and the probability of a given stimulus appearing as a test item, it produced the fastest response latencies.

2. The slope, representing the increase in reaction time with an increased number of items in the Theoretical Positive Set, was shallowest in the Sounds Condition of Experiment I. If we accept the proposition that these increments reflect the operation of a memory-related stage of processing, we must conclude that attention to the auditory features of

the to-be-memorized items facilitates such processing, relative to emphasizing the visual or semantic features of the stimuli. This conclusion is essentially independent of the assumptions of Sternberg's serial exhaustive search model. That is, the shallow slope of the Sounds Condition, relative to the other two conditions, is evidence for auditory memory codes whether the memory search is serial or parallel, exhaustive or self-terminating. The only necessary assumption for this conclusion is that the increments in reaction time with larger Positive Sets reflect the use of memory codes, rather than some other stage of processing.

3. There was both less total variability and less intra-set variability within the Sounds Condition of Experiment I than in the other two conditions. Therefore, the opportunity to organize the to-be-memorized items according to their auditory properties provides the most efficient basis for the memory-dependent response classification. The Sounds Condition was again at the greatest disadvantage since it involved the greatest discrepancy between the set of displayed stimuli and the set of potentially correct items. That is, the Sounds Condition provided the least stimulus (D variable) to target (C variable) compatibility. During the Looks instructions, the subject simply gave a Yes response to the same items as in the displayed set on all 48 trials. By comparison, only 4 of the 48 trials permitted this direct stimulus to target correspondence in the Sounds Condition. For the remaining 44 trials the subject had to determine which additional (non-displayed) items should also receive a Yes response. Despite this added complexity, the subjects gave their most consistent performance under the Sounds instruction.

4. The analysis of errors revealed an overwhelming tendency to incorrectly respond Yes to a homonym of a Positive Set item in the Looks and Means Conditions. By contrast, there was no substantial evidence of visual or semantic confusions. The evidence from false positive responses alone was so strong that it warrants the conclusion that memory codes retain auditory properties. Therefore, we need not rely solely upon latency measures to verify that memory is auditory.

5. The slope of the Sounds Condition was independent of the type of stimulus materials, within the range tested by the present research. This includes both with and without homonyms present and for numbers as well as three-letter words (Sounds 1:3 vs. Sounds 1:1). This result contradicts the claim that the stimulus probability is the critical factor. Stimulus probability varied over a 3:1 range in comparing the Sounds 1:3 with the Sounds 1:1 Conditions, since these conditions represented a 3:1 vs. a 1:1 ratio of physically distinct homonym versions per sound. This result is most reasonable if a serial exhaustive search model is assumed, but it is equally sound for other models, provided that the memory code is auditory.

6. The same slope was obtained with instructions emphasizing other than auditory features of the stimuli, according to the Sound-Tag interpretation. This finding includes both results of Experiment I (Looks and Means Conditions) and of Experiment III (Looks without homonyms present). This result adds to our confidence that the theoretical basis for this prediction is well founded (i.e., that subjects engage in a serial exhaustive search through auditory codes). Even with a task designed to favor some other encoding strategy (visual or semantic memory), subjects persist in using auditory codes -- despite the fact that the auditory codes lead to homonym errors.

A possible objection to the conclusion that short-term recognition memory involves solely auditory codes is that memory is multidimensional. This position has gained support from experiments in which it has been shown that subjects can respond to auditory, visual, or semantic features, as directed. On this basis alone, the multidimensional position is indistinguishable from the present position which holds that the primary mode of encoding is auditory with tags appropriate to other features attached to the name codes. While stimulus classifications made briefly after the onset of a to-be-memorized item may benefit from visual similarity, the present study has only found evidence of auditory properties for time intervals ranging from 5 to 23 seconds after the onset of the to-be-memorized list. Within this interval, we have discovered memory-related processes which require under 25 milliseconds per comparison. To consider that semantic features are available some 200-300 msec. later than auditory

properties (as in Shulman's results, 1970), permits a large amount of time for other cognitive operations to take place on the auditory code. We have made the strong assumption that elapsed time implies elapsed processing, an assumption which has not been disconfirmed in the present findings.

While the present findings are not inconsistent with the multidimensional trace models, they do not provide evidence for any other properties than auditory ones. The results are compatible with the unitary code theory, a position which makes stronger (i.e., easier to disconfirm) predictions. It is not clear what evidence could disconfirm the more general versions of a multidimensional theory.

In tempering the conclusion that memory codes are auditory we should note once more that the present research used only verbal materials, and that we are unable to state whether this acoustic-like memory code follows more closely the perception of speech, the articulation of words, or the physical properties of sound waves.

While the conclusion that memory is auditory for verbal materials does not rest on any single fact, nor on the frailty of the totality of facts, we are compelled to accept this conclusion given the convergence of evidence from diverse sources. The conclusion relies on neither a single dependent variable (absolute latency, slopes, and errors have been considered), nor on one set of theoretical assumptions (not just the serial exhaustive search model). Nevertheless, the entire set of findings form a consistent picture when interpreted according to the serial exhaustive search model proposed by Sternberg (1969). It is this model which permits the straight-forward interpretation of the magnitude of the slope (i.e., about 23 msec. per sound) as representing the time to make a single memory comparison. Furthermore, this model should be credited with both prompting these experiments and parsimoniously organizing their results.

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APPENDICES

APPENDIX F

ILLUSTRATED INSTRUCTIONS FOR EXPERIMENT II

"The Same Set"

I	2	4	(arabic)
I	II	IV	(roman)
ONE	TWO	FOUR	(english #)
WON	TOO	FOR	(english words)

Example:

2 FOR	I	VIN	TOO	IV
	yes	no	yes	no

